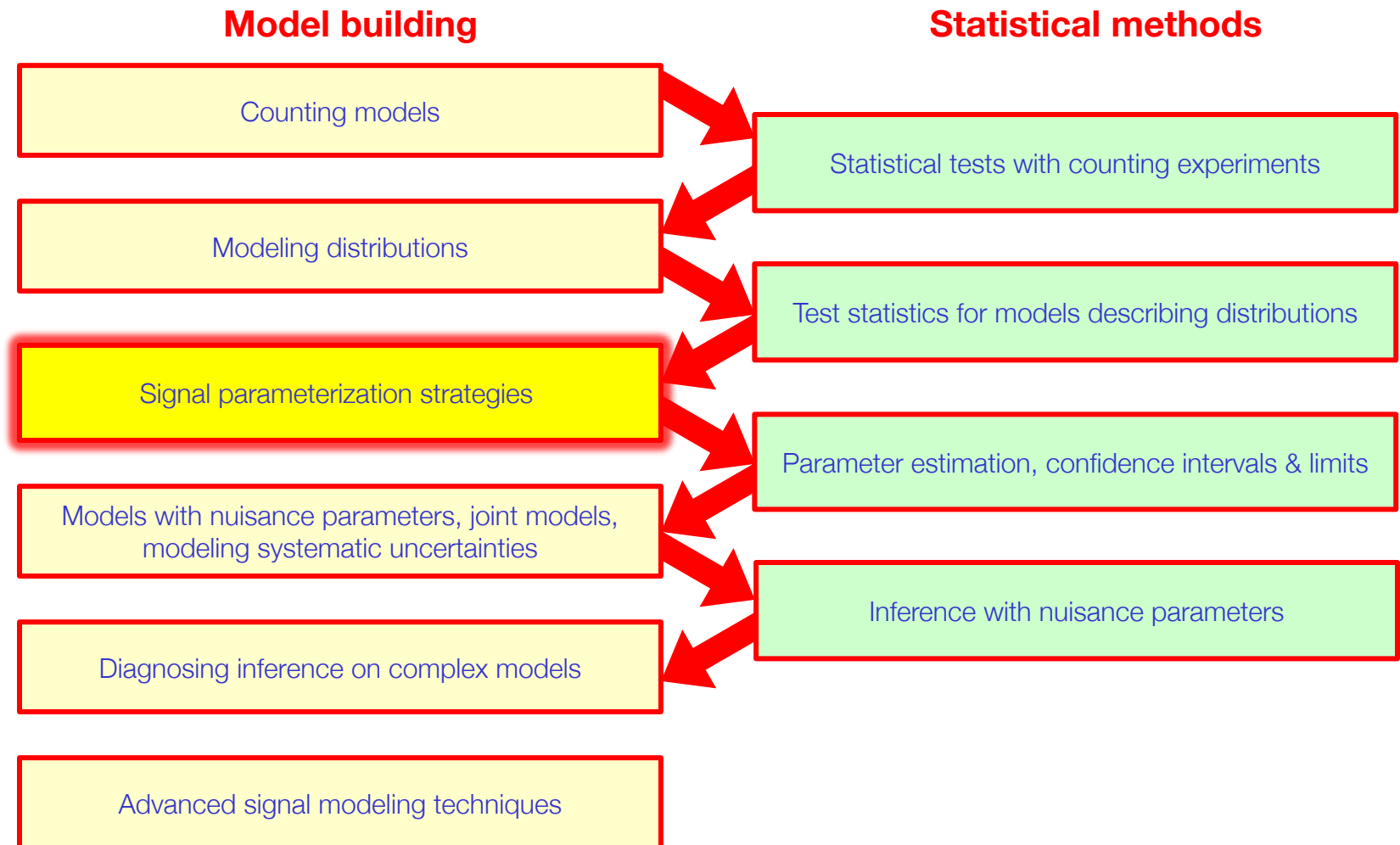


# Model building 3

Models with parameters I -  
analytical parametric models,  
multi-dimensional models  
template morphing approach for  
histogram-based models

# Roadmap of this course

- Start with basics, gradually build up to complexity

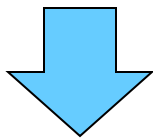


## Introduce concept of composite hypotheses

- In most cases in physics, a hypothesis is not “simple”, but “composite”
- **Composite hypothesis** = Any hypothesis which does *not* specify the population distribution completely
- Example: counting experiment with signal and background, that leaves signal expectation unspecified

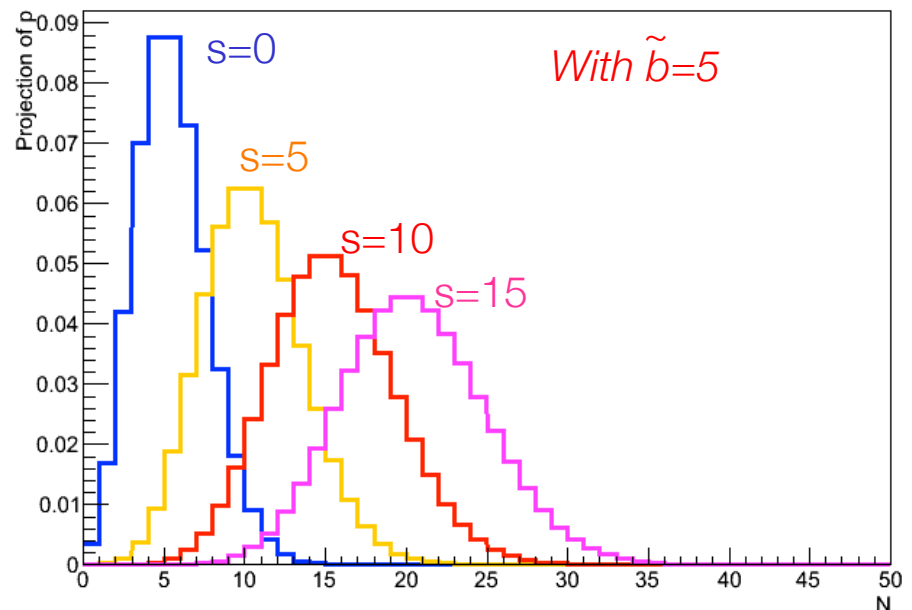
Simple hypothesis

$$L = \text{Poisson}(N | \tilde{s} + \tilde{b})$$



$$L(s) = \text{Poisson}(N | s + \tilde{b})$$

Composite hypothesis



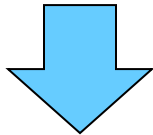
(My) notation convention: all symbols with  $\sim$  are constants

# A common convention in the meaning of model parameters

- A common convention is to recast signal rate parameters into a normalized form (e.g. w.r.t the Standard Model rate)

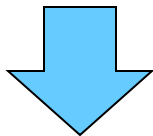
Simple hypothesis

$$L = \text{Poisson}(N \mid \tilde{s} + \tilde{b})$$



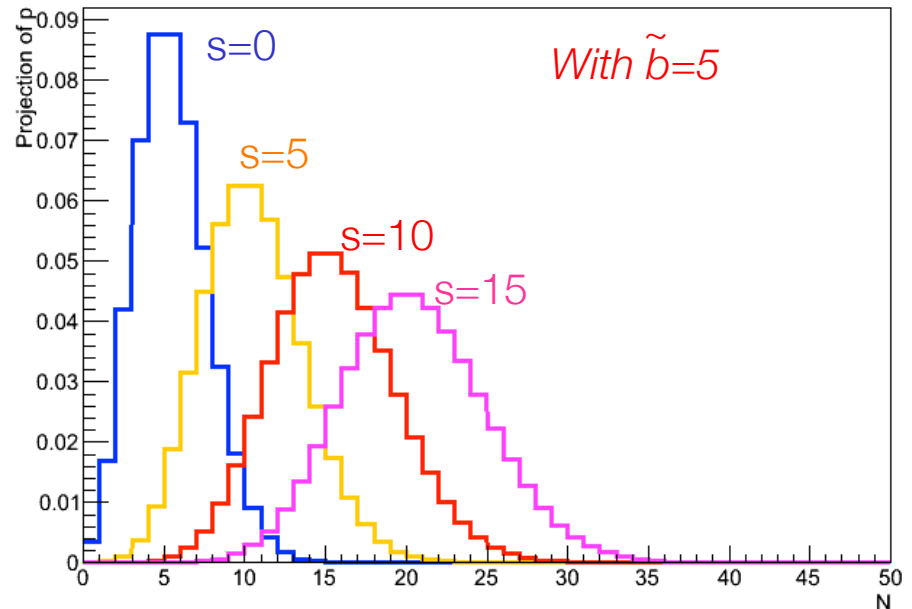
$$L(s) = \text{Poisson}(N \mid s + \tilde{b})$$

Composite hypothesis



$$L(\mu) = \text{Poisson}(N \mid \mu \cdot \tilde{s} + \tilde{b})$$

Composite hypothesis  
with normalized rate parameter



*'Universal' parameter interpretation  
makes it easier to work with your models*

$\mu=0 \rightarrow$  no signal

$\mu=1 \rightarrow$  expected signal

$\mu>1 \rightarrow$  more than expected signal

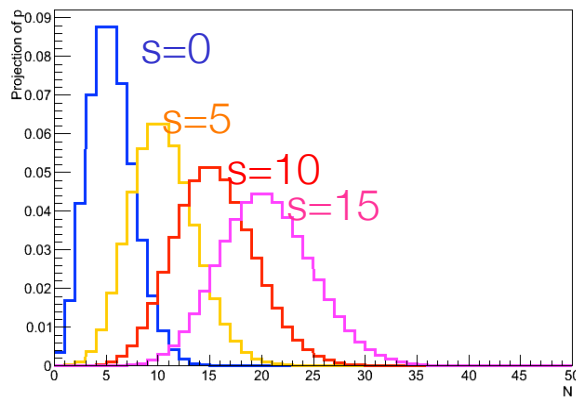
# What can we do with composite hypothesis

- With simple hypotheses – inference is restricted to making statements about  $P(D|hypo)$  or  $P(hypo|D)$
- With composite hypotheses – many more options
- **1 Parameter estimation and variance estimation**
  - What is value of  $s$  for which the observed data is most probable?
  - What is the variance (std deviation squared) in the estimate of  $s$ ? }  $s=5.5 \pm 1.3$
- **2 Confidence intervals**
  - Statements about model parameters using frequentist concept of probability
  - $s < 12.7$  at 95% confidence level
  - $4.5 < s < 6.8$  at 68% confidence level
- **3 Bayesian credible intervals**
  - Bayesian statements about model parameters
  - $s < 12.7$  at 95% credibility

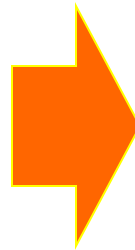
# Model building for discovery, X-section $\rightarrow$ yield parameter

0-dimensional (counting)

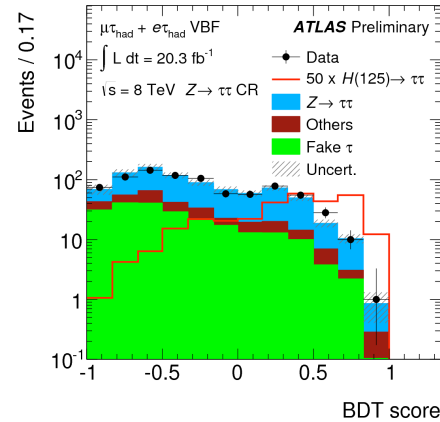
1-dimensional (discriminant)



$$\text{Poisson}(N|\mathbf{S}+B)$$

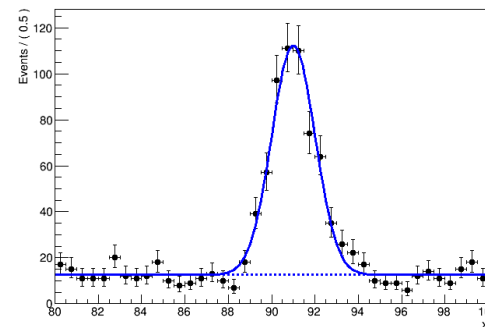


MVA discriminant



$$\mathbf{S}^* \text{sig}(x) + \mathbf{B}^* \text{bkg}(x)$$

Physics-inspired discriminant



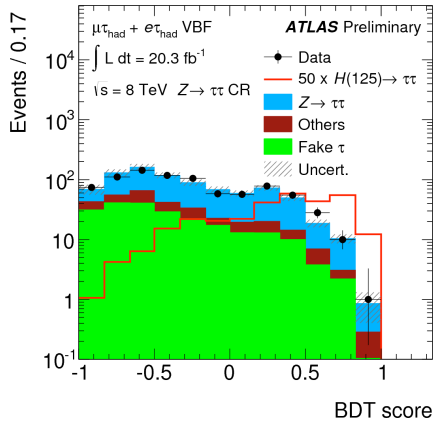
$$\mathbf{S}^* \text{sig}(x) + \mathbf{B}^* \text{bkg}(x)$$

# Models for discovery, X-section $\rightarrow$ yield parameter

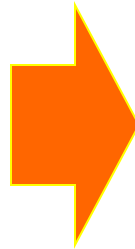
1-dimensional (discriminant)

2-dimensional?

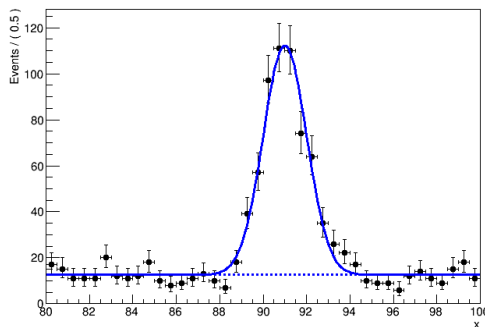
MVA discriminant



$$S^* \text{sig}(x) + B^* \text{bkg}(x)$$



Physics-inspired discriminant



$$S^* \text{sig}(x) + B^* \text{bkg}(x)$$

**Q: When is it useful to build probability models in  $\geq 2$  observables?**

A1: When you have a physics model with a clear prediction for the full 2D model..

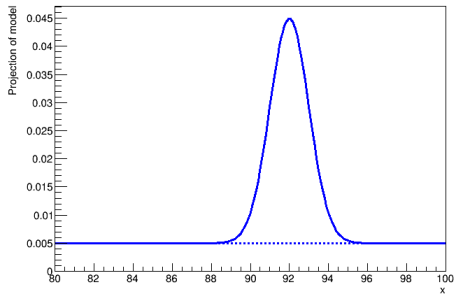
Often you don't and then you let an MVA reduce the n-Dim space to 1-dimension

But sometimes you have clear models described 2 or more observables  $\rightarrow$  No point in letting an MVA approximate what you know analytically.

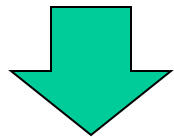
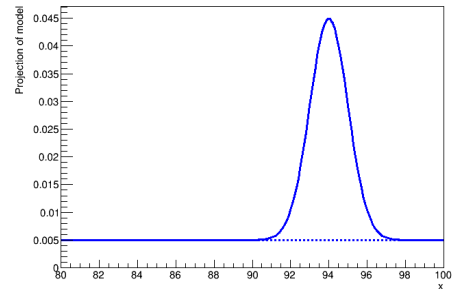
# Case study – dependence of 1-D model on another observable

- A common scenario for 2D modelling is the following: You observe that the mean reconstructed mass of some particle depends on another observable

Model for mass at (y=0)  
 $\text{sig}(m) = \text{Gaussian}(m, 92, 1)$

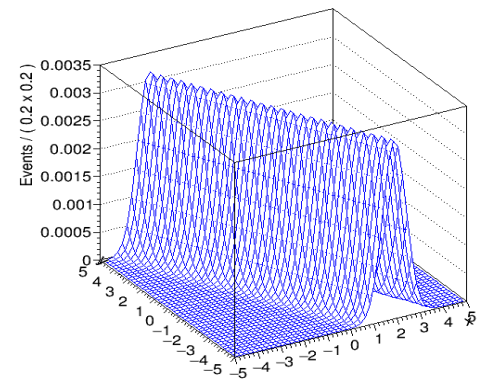


Model for mass at (y=3)  
 $\text{sig}(m) = \text{Gaussian}(m, 94, 1)$



$$\text{sig}(m, y) = \text{Gaussian}(m, \text{mean}(y), 1)$$

Solution:  
 introduce a  
 function **mean(y)**  
 that describes  
 dependence  
 of mean of y



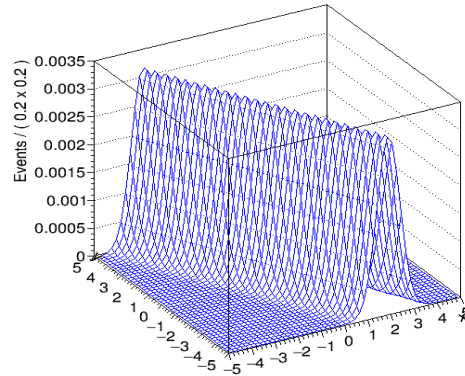
Q: Is sig(m,y) a proper 2-dimensional model?



# Case study – dependence of 1-D model on another observable

$$\text{sig}(m,y) = \text{Gaussian}(m, \text{mean}(\mathbf{y}), 1)$$

Solution:  
introduce a  
function **mean(y)**  
that describes  
dependence  
of mean of y



Q: Is sig(m,y) a proper  
2-dimensional model?



A: No!  
Distribution in y is  
unlikely to be flat...

- Challenge for 2D models: distributions in x,y and all correlations must all be correct! Seems intractable, but solutions exists
- Instead of immediately defining a 2D model **f(x,y)**, define first the *conditional* probability density function **f(x|y)**

$$\begin{aligned} f(x,y) \\ = \\ \text{2D model for} \\ \text{both x and y} \\ \int f(x,y) dx dy \equiv 1 \end{aligned}$$

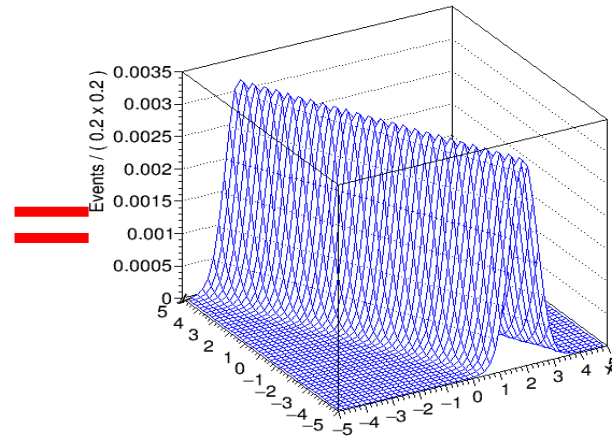
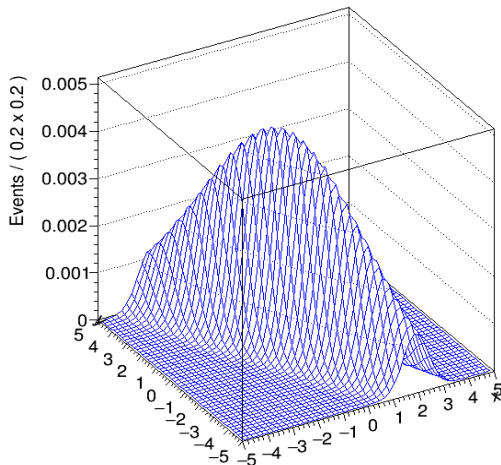
$$\begin{aligned} f(x|y) \\ = \\ \text{1D model for x} \\ \text{at a given value of y} \\ \int f(x,y) dx \equiv 1 \quad \forall y \end{aligned}$$

*This is really what  
we meant when we  
formulated this:  
Gaussian(m, mean(y), 1)*

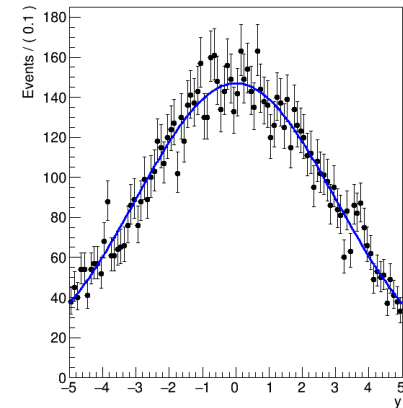
# Case study – dependence of 1-D model on another observable

- Given a conditional model  $f(x|y)$  can build full 2D model by multiplying with a model  $g(y)$

$$\text{sig}(m,y) = \text{sig}_m(m|y) * \text{sig}_y(y)$$



\*

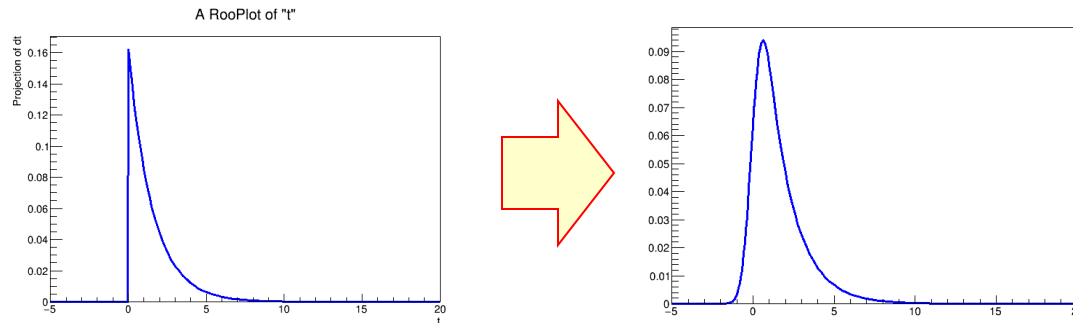
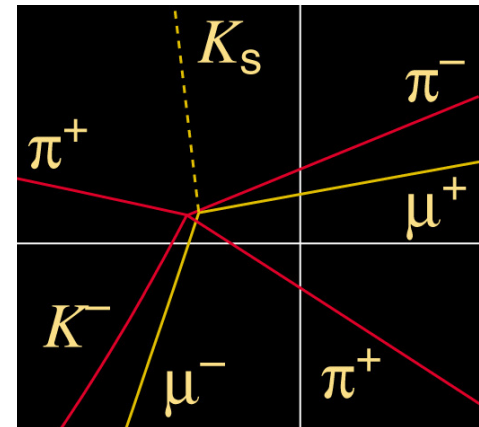


Gaussian( $m$ , **mean(y)**, 1)

Gaussian(y)

## Case study – per-event errors

- Another common variant of this type of modeling problem is the so-called ‘per-event’ error
- Example: observable = decay time distribution, measured from reconstructed vertex.
  - In absence of a detector resolution, exponential decay distribution
  - In real life, distribution is convoluted with (Gaussian) reconstruction resolution



- But vertex reconstruction gives also estimate of uncertainty for every reconstructed vertex → the ‘per-event error’
  - Can take this into account: well-reconstructed events carry more information
- How? Scale assumed resolution with per-event error

$$f(t | \delta t) = \text{Decay}(t) \otimes \text{Gaussian}(t, 0, \sigma \cdot \delta t)$$

# Case study – per-event errors

- Visualization of decay function with variable resolution

Decay function (symmetrized)  
convoluted with Gaussian resolution  
at 4 different values of per-event error

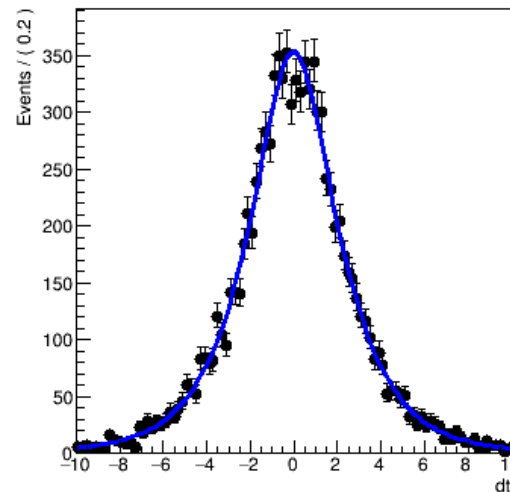
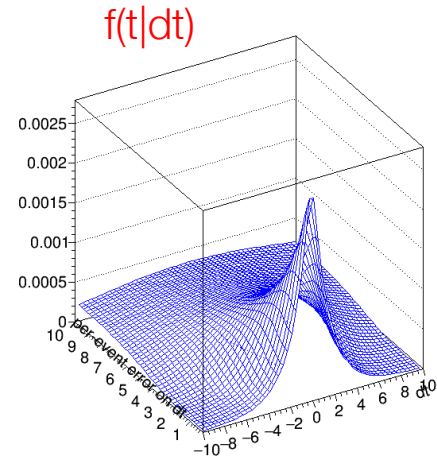
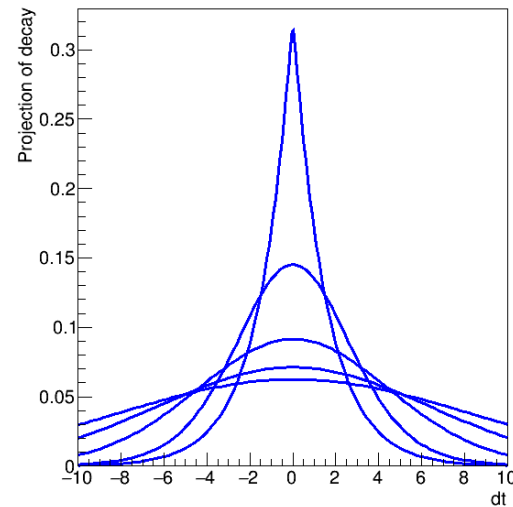
$$f(t | \delta t) = \text{Decay}(t) \otimes \text{Gaussian}(t, 0, \sigma \cdot \delta t)$$

Gain: high-resolution events  
carry more weight in likelihood →  
better estimate of model parameters

Full 2D-model:  
 $F(t, dt) = F_1(t|dt) * F_2(dt)$

Shown here: *projection on t*  
 $F(t) = \text{Int} [ F_1(t|dt) * F_2(dt) ] dt$

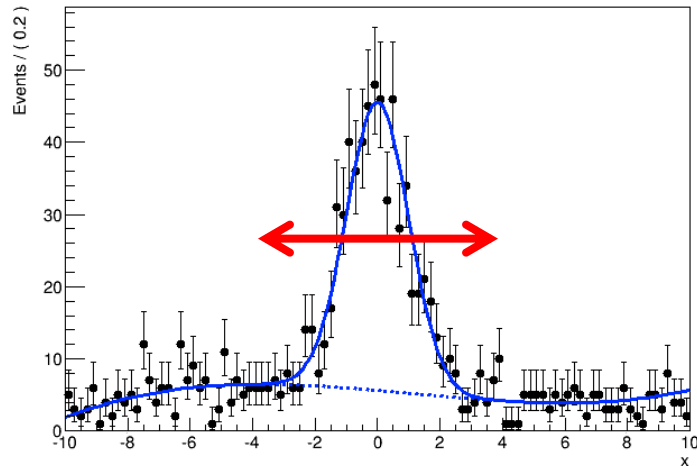
Slices of decay(dt|dt\_err) at various dt\_err



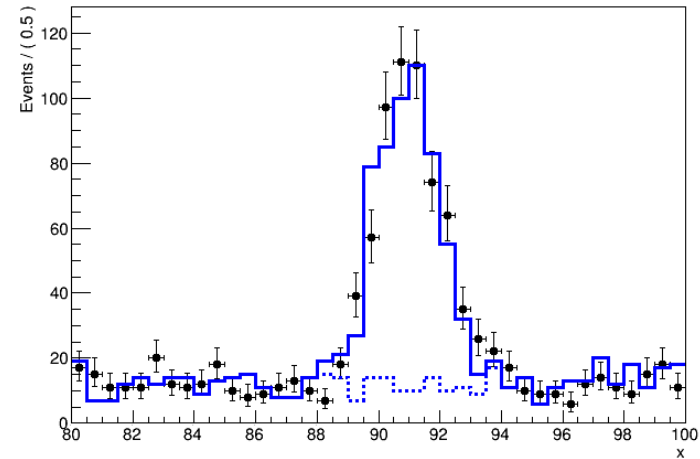
# Model building for measurements → shape parameter

- Beyond discovery/rate measurements, can also build models to measure properties of particles (e.g mass)  
→ introduce shape parameters
- Often trivial for analytical models,  
less so for simulation-based models

$$F(x|\mathbf{m}) = \text{Gaussian}(x, \mathbf{m}, \sigma) + \text{bkg}$$

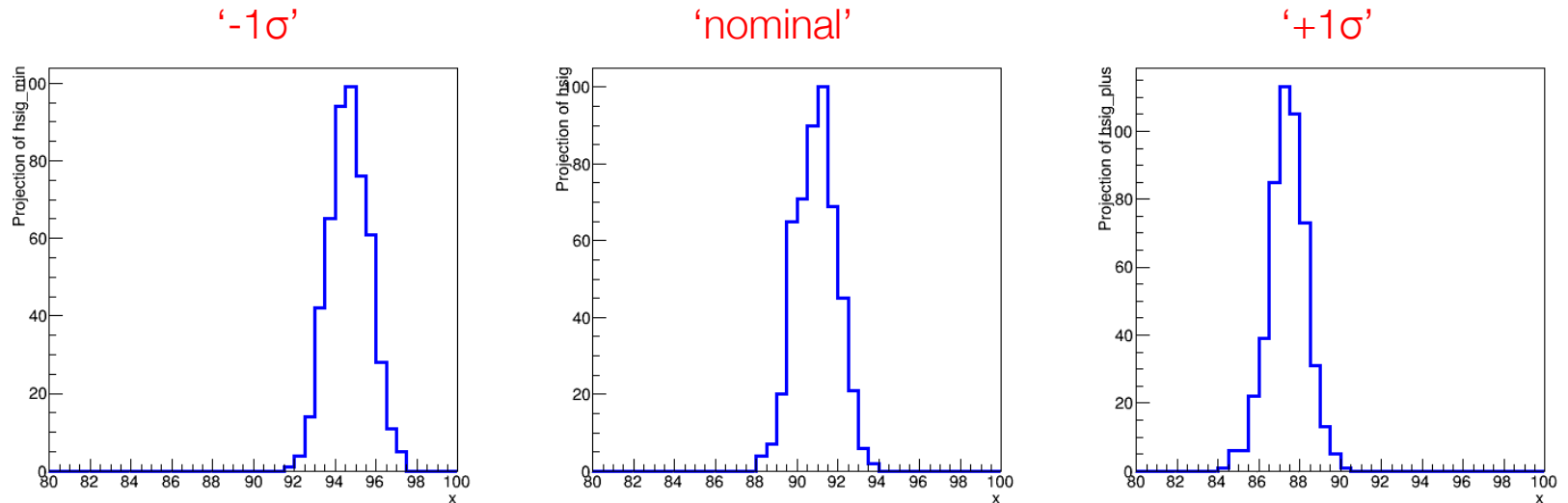


$$F(x|\mathbf{m}) = ??$$



# Modeling of shape variations in the likelihood

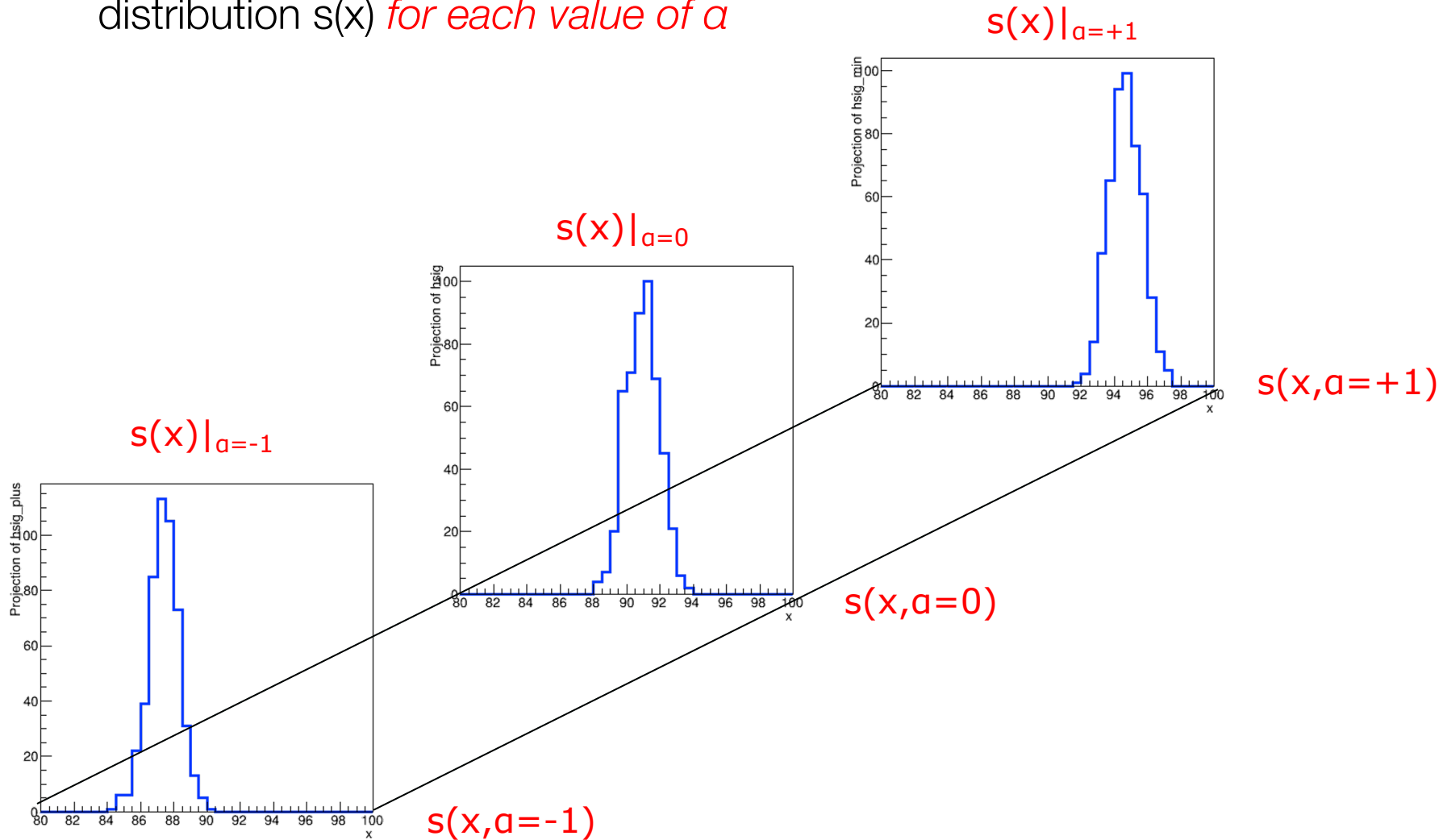
- If underlying simulation has free parameter  $\theta$ , can assess impact on reconstructed shapes by rerunning simulation at different values
  - Obtain histogram templates for distributions at '+1 $\sigma$ ' and '-1 $\sigma$ ' settings of systematic effect



- Challenge: **construct an empirical response function based on the interpolation of the shapes of these three templates.**

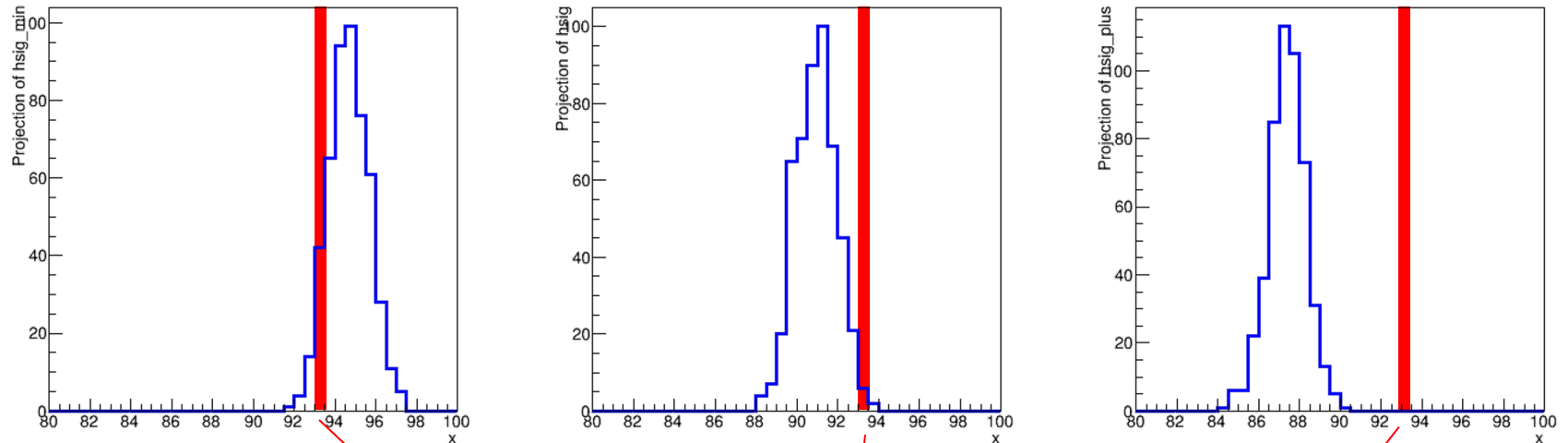
# Need to interpolate between template models

- Need to define ‘morphing’ algorithm to define distribution  $s(x)$  *for each value of  $a$*

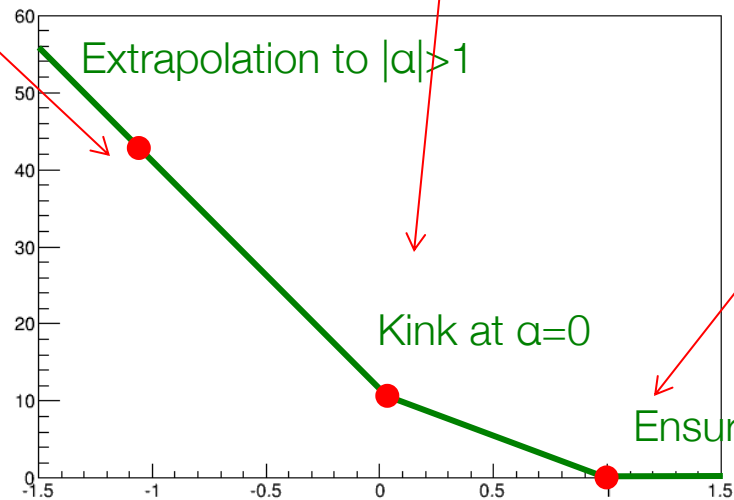


# Piecewise linear interpolation

- Simplest solution is piece-wise linear interpolation for each bin

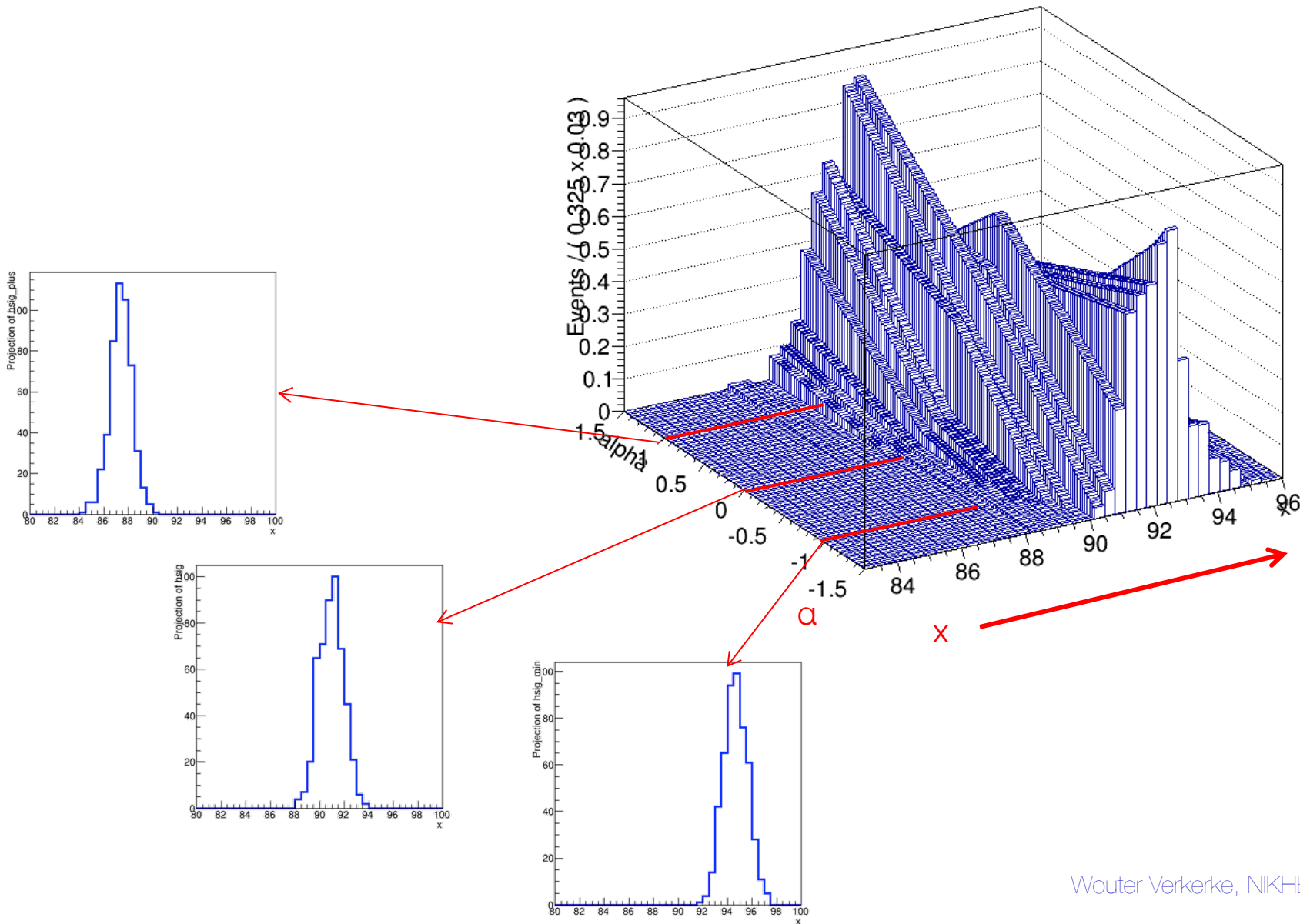


Piecewise linear interpolation response model for a one bin



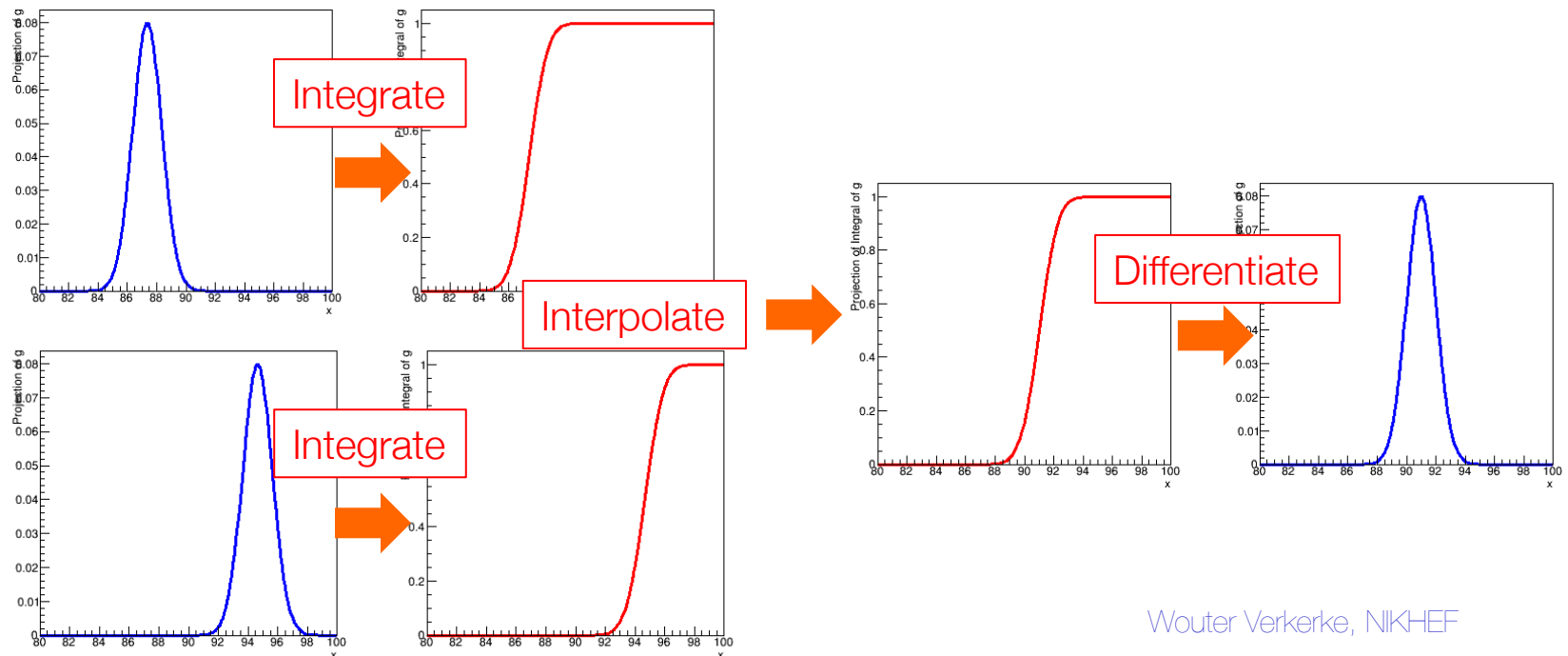


# Visualization of bin-by-bin linear interpolation of distribution



# Other morphing strategies – ‘horizontal morphing’

- Other template morphing strategies exist that are less prone to unintended side effects
- A ‘horizontal morphing’ strategy was invented by Alex Read.
  - Interpolates the cumulative distribution function instead of the distribution
  - Especially suitable for shifting distributions
  - Here shown on a continuous distribution, but also works on histograms
  - Drawback: computationally expensive, algorithm only worked out for 1 NP



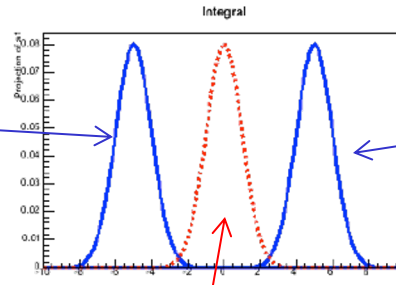
# Yet another morphing strategy – ‘Moment morphing’

*M. Baak & S. Gadatsch*

- Given two template model  $f_-(x)$  and  $f_+(x)$  the strategy of moment morphing considers first two moment of template models (mean and variance)

$$\mu_- = \int x \cdot f_-(x) dx$$

$$V_- = \int (x - \mu_-)^2 \cdot f_-(x) dx$$



$$\mu_+ = \int x \cdot f_+(x) dx$$

$$V_+ = \int (x - \mu_+)^2 \cdot f_+(x) dx$$

- The goal of moment morphing is to construct an interpolated function that has linearly interpolated moments

$$\begin{aligned} \mu(\alpha) &= \alpha\mu_- + (1 - \alpha)\mu_+ \\ V(\alpha) &= \alpha V_- + (1 - \alpha)V_+ \end{aligned} \quad [1]$$

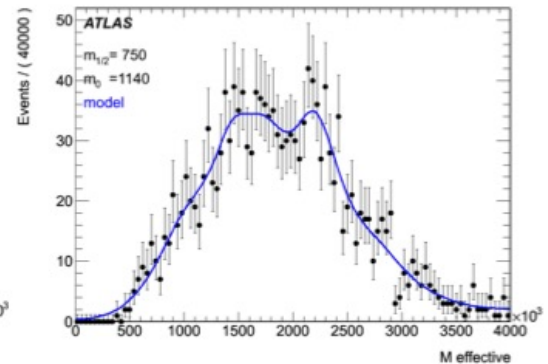
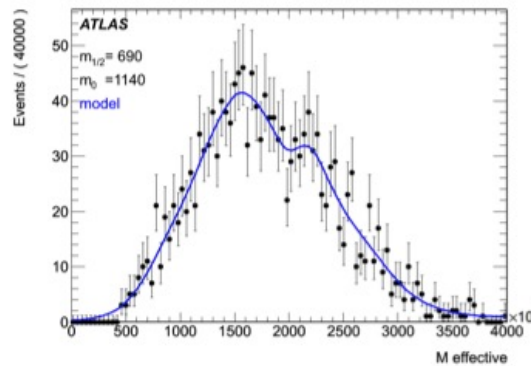
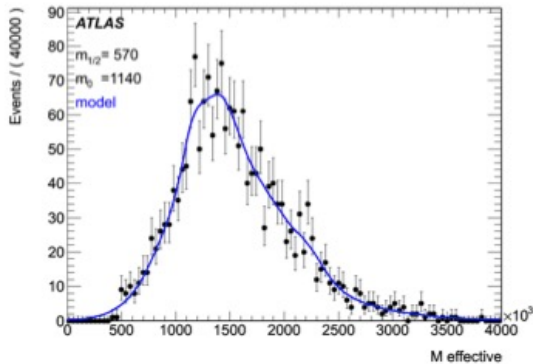
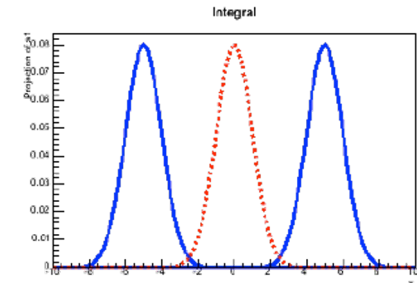
- It constructs this morphed function as combination of linearly transformed input models

$$f(x, \alpha) \rightarrow \alpha f_-(ax + b) + (1 - \alpha) f_+(cx - d)$$

- Where constants a,b,c,d are chosen such so that  $f(x, \alpha)$  satisfies conditions [1]

# Yet another morphing strategy – ‘Moment morphing’

- For a Gaussian probability model with linearly changing mean and width, moment morphing of two Gaussian templates is the exact solution
- But also works well on ‘difficult’ distributions



- Good computational performance
  - Calculation of moments of templates is expensive, but just needs to be done once, otherwise very fast (just linear algebra)

$$f(x, \alpha) \rightarrow \alpha f_-(ax + b) + (1 - \alpha) f_+(cx - d)$$

- Multi-dimensional interpolation strategies exist

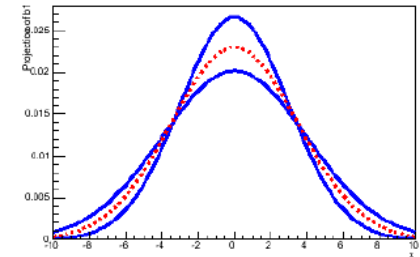
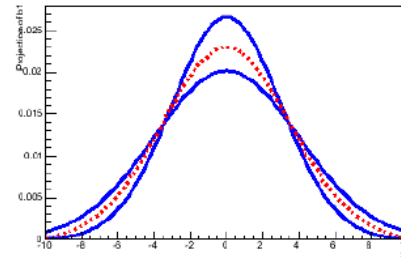
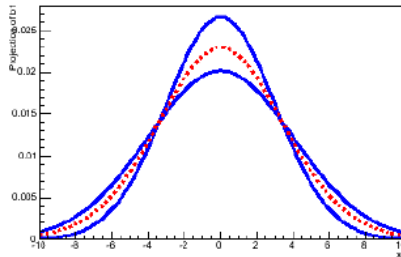
# There are other morphing algorithms to choose from

Vertical Morphing

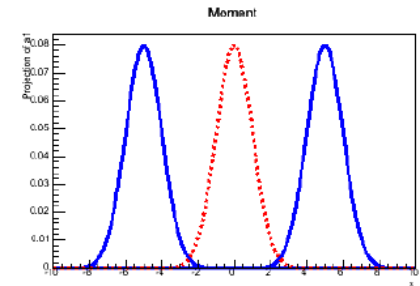
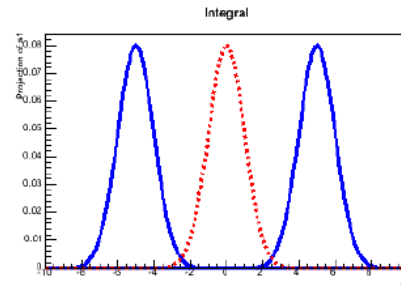
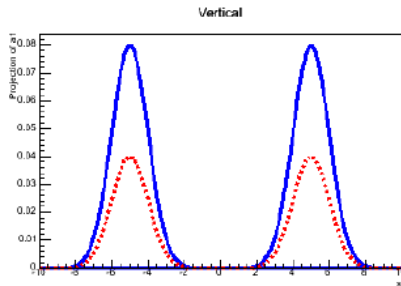
Horizontal Morphing

Moment Morphing

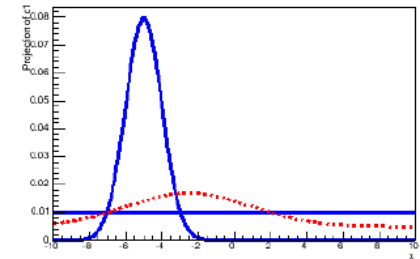
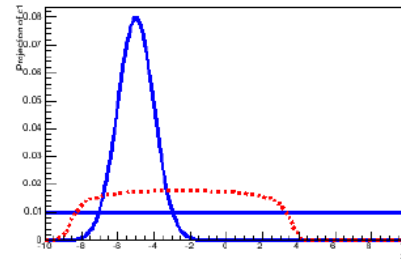
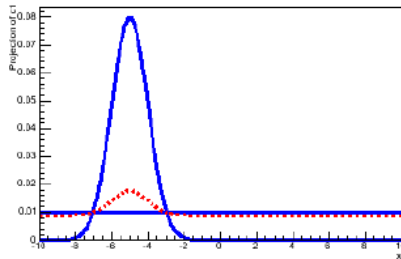
Gaussian varying width



Gaussian varying mean



Gaussian to Uniform (this is conceptually ambiguous!)



n-dimensional morphing?

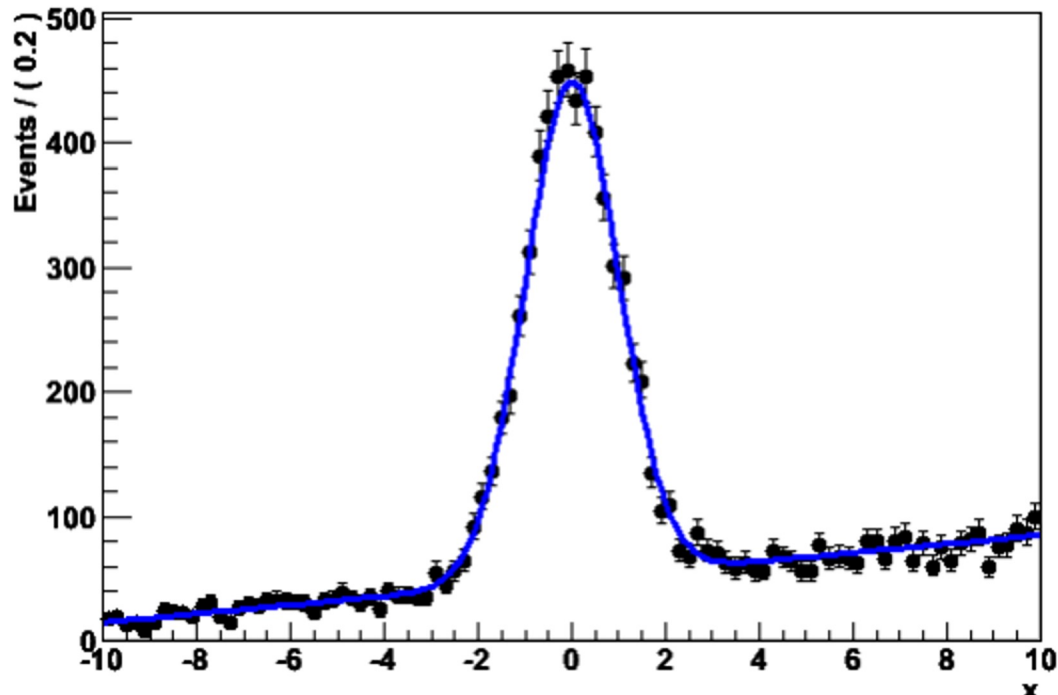


# Software tools 1

Basic RooFit modeling

## Roofit – Focus: coding likelihood functions

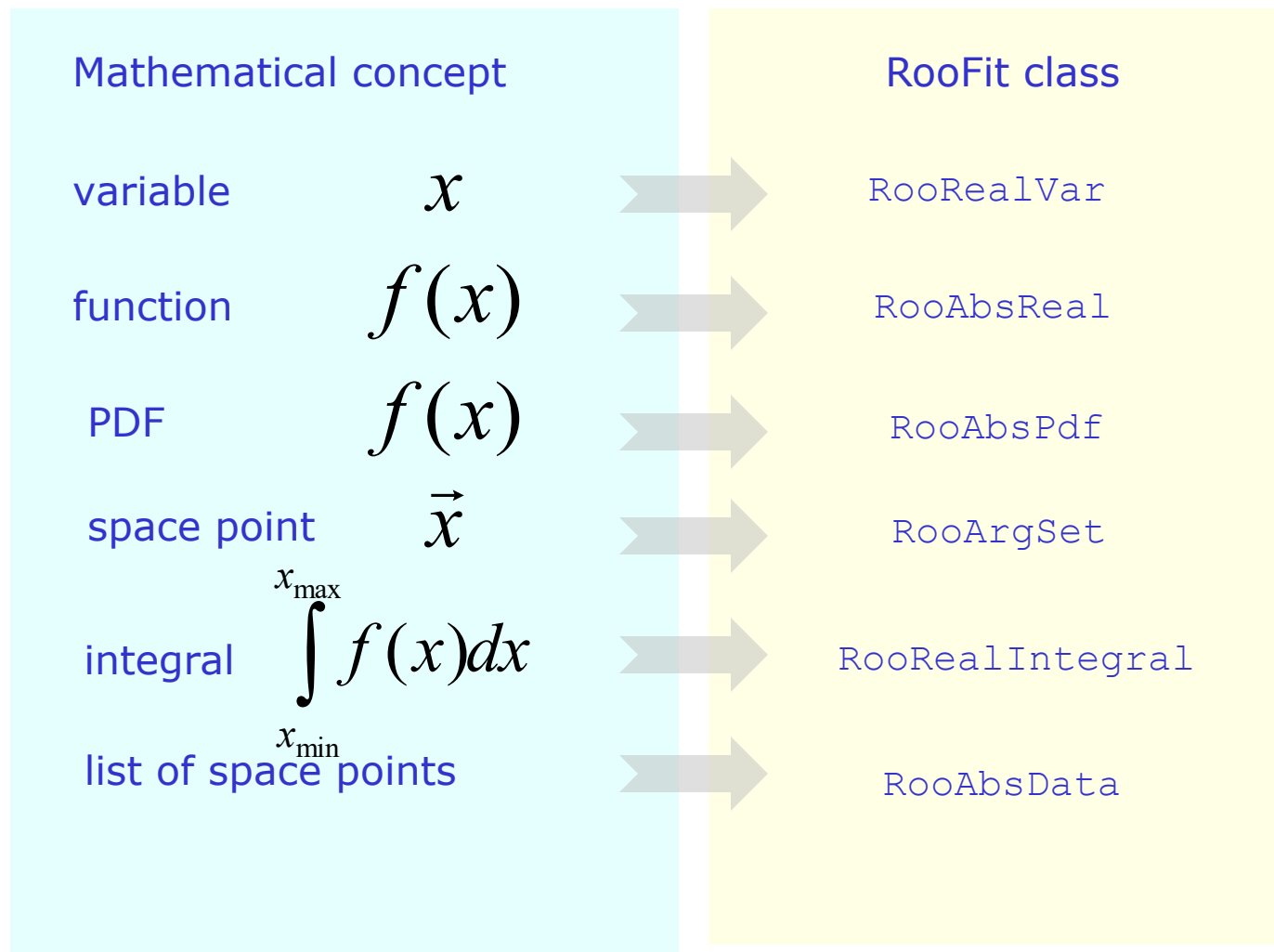
- Focus on one practical aspect of many data analysis in HEP: **How do you formulate your likelihood functions in ROOT**
  - For ‘simple’ problems (gauss, polynomial) this is easy



- But if you want to do unbinned ML fits, use non-trivial functions, or work with multidimensional functions you quickly find that you need some tools to help you

# RooFit core design philosophy

- Mathematical objects are represented as C++ objects





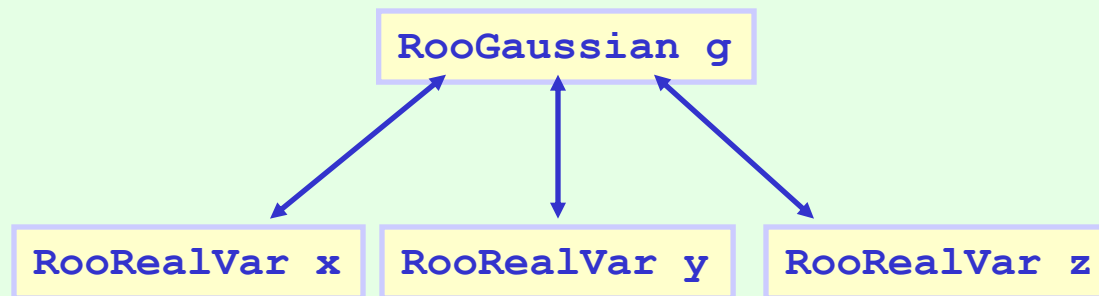
# RooFit core design philosophy - Workspace

- Instead of `double Likelihood(double paramVec[])`, a flexible modular structure of 'programmed' functions

Math

Gauss( $x, \mu, \sigma$ )

RooFit diagram



RooFit code

```
RooRealVar x("x","x",-10,10) ;  
RooRealVar m("m","y",0,-10,10) ;  
RooRealVar s("s","z",3,0.1,10) ;  
RooGaussian g("g","g",x,m,s) ;
```

# Basics – Creating and plotting a Gaussian p.d.f

Setup gaussian PDF and plot

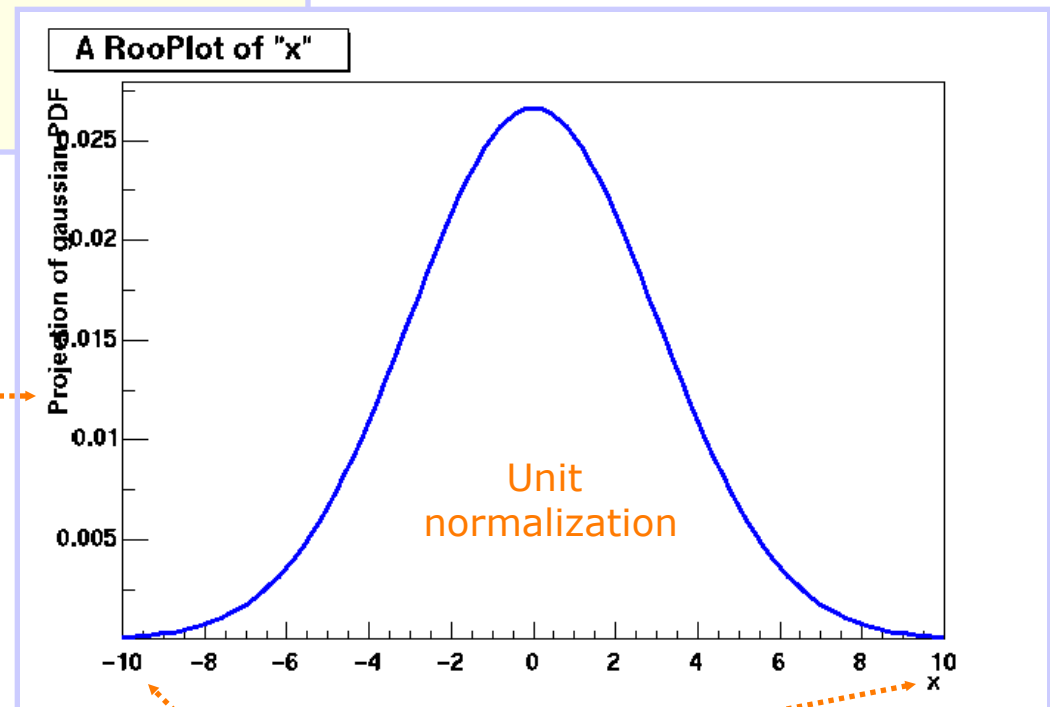
```
// Create an empty plot frame
RooPlot* xframe = w::x.frame() ;

// Plot model on frame
model.plotOn(xframe) ;

// Draw frame on canvas
xframe->Draw() ;
```

Axis label from gauss title

A RooPlot is an empty frame capable of holding anything plotted versus it variable



Plot range taken from limits of x

# Basics – Generating toy MC events

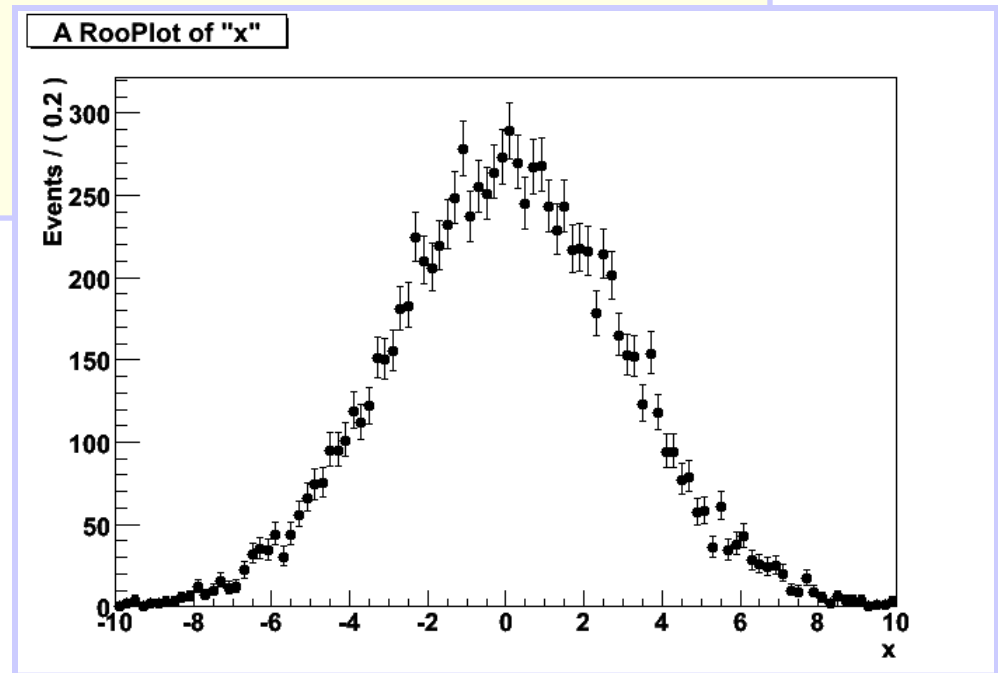
Generate 10000 events from Gaussian p.d.f and show distribution

```
// Generate an unbinned toy MC set
RooDataSet* data = w::gauss.generate(w::x,10000) ;

// Generate an binned toy MC set
RooDataHist* data = w::gauss.generateBinned(w::x,10000) ;

// Plot PDF
RooPlot* xframe = w::x.frame()
data->plotOn(xframe) ;
xframe->Draw() ;
```

Can generate both binned and unbinned datasets

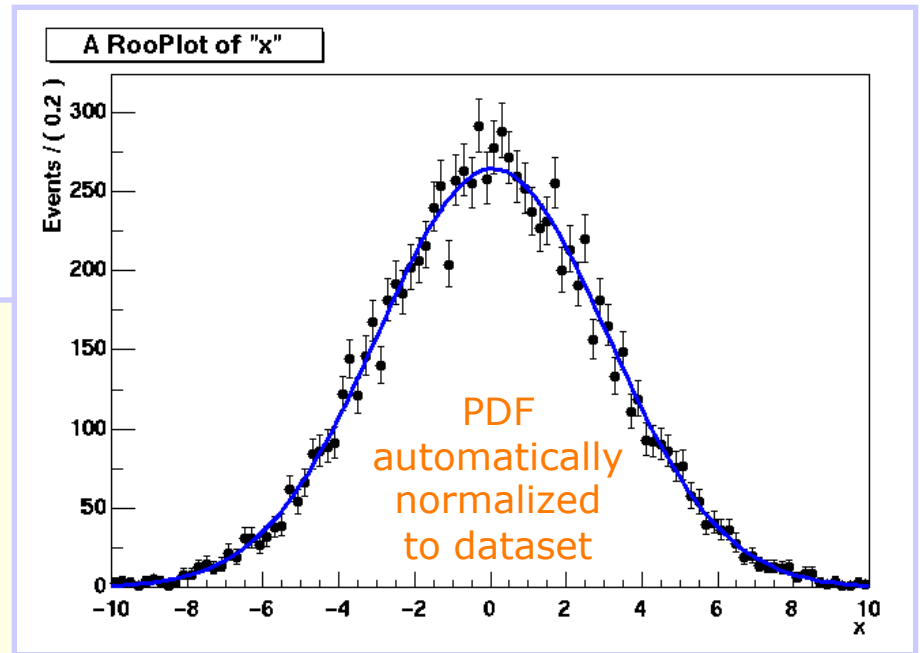


# Basics – ML fit of p.d.f to *unbinned* data

```
// ML fit of gauss to data
w::gauss.fitTo(*data) ;
(MINUIT printout omitted)

// Parameters if gauss now
// reflect fitted values
w::mean.Print()
RooRealVar::mean = 0.0172335 +/- 0.0299542
w::sigma.Print()
RooRealVar::sigma = 2.98094 +/- 0.0217306

// Plot fitted PDF and toy data overlaid
RooPlot* xframe = w::x.frame() ;
data->plotOn(xframe) ;
w::gauss.plotOn(xframe) ;
```



# RooFit core design philosophy - Workspace

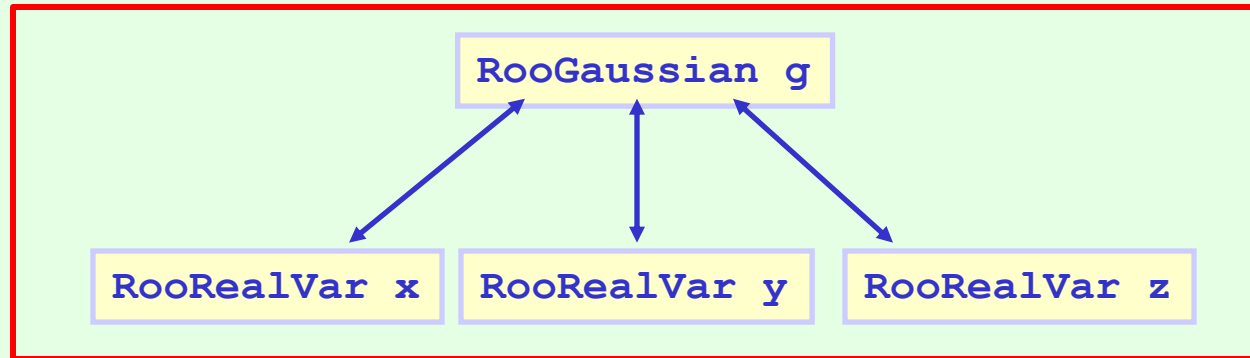
- The workspace serves a container class for all objects created

Math

Gauss( $x, \mu, \sigma$ )

RooWorkspace

RooFit diagram

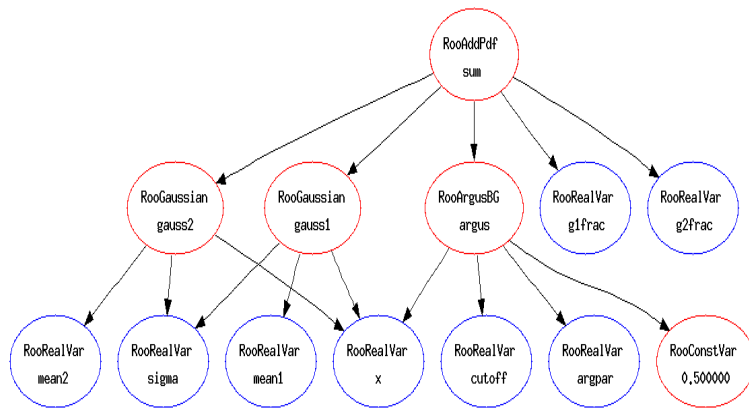


RooFit code

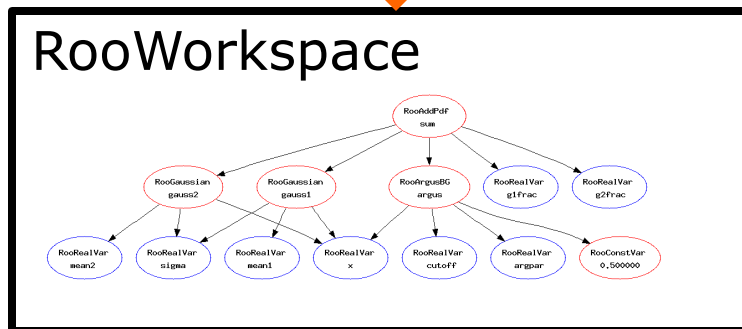
```
RooRealVar x("x","x",-10,10) ;  
RooRealVar m("m","y",0,-10,10) ;  
RooRealVar s("s","z",3,0.1,10) ;  
RooGaussian g("g","g",x,m,s) ;  
RooWorkspace w("w") ;  
w.import(g) ;
```

# The workspace

- The workspace concept has revolutionized the way people share and combine analysis
  - **Completely** factorizes process of building and using likelihood functions
  - You can give somebody an analytical likelihood of a (potentially very complex) physics analysis in a way to the easy-to-use, provides introspection, and is easy to modify.



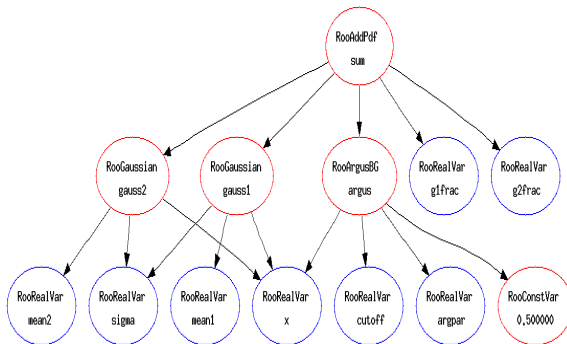
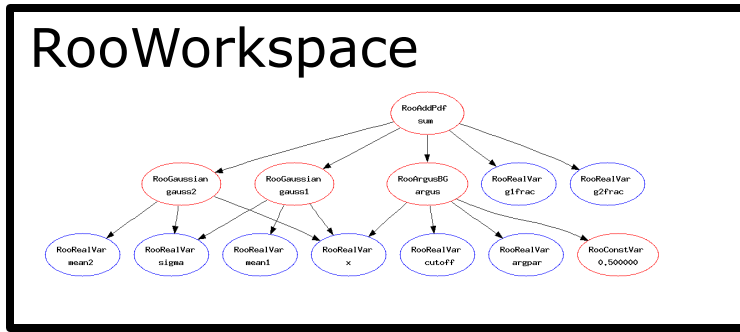
```
RooWorkspace w("w") ;  
w.import(sum) ;  
w.writeToFile("model.root") ;
```



model.root



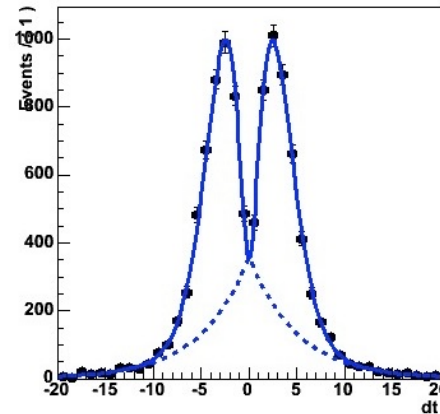
# Using a workspace



```
// Resurrect model and data
TFile f("model.root") ;
RootWorkspace* w = f.Get("w") ;
RootAbsPdf* model = w->pdf("sum") ;
RootAbsData* data = w->data("xxx") ;
```

```
// Use model and data
model->fitTo(*data) ;
```

```
RootPlot* frame =
    w->var("dt")->frame() ;
data->plotOn(frame) ;
model->plotOn(frame) ;
```



# Factory and Workspace

- *One C++ object per math symbol* provides ultimate level of control over each objects functionality, but results in lengthy user code for even simple macros
- Solution: add factory that auto-generates objects from a math-like language. **Accessed through factory() method of workspace**
- Example: reduce construction of Gaussian pdf and its parameters from 4 to 1 line of code

```
RooRealVar x("x","x",-10,10) ;  
RooRealVar mean("mean","mean",5) ;  
RooRealVar sigma("sigma","sigma",3) ;  
RooGaussian f("f","f",x,mean,sigma) ;  
w.import(f) ;
```



```
w.factory("Gaussian::f(x[-10,10],mean[5],sigma[3])") ;
```



# RooFit core design philosophy - Workspace

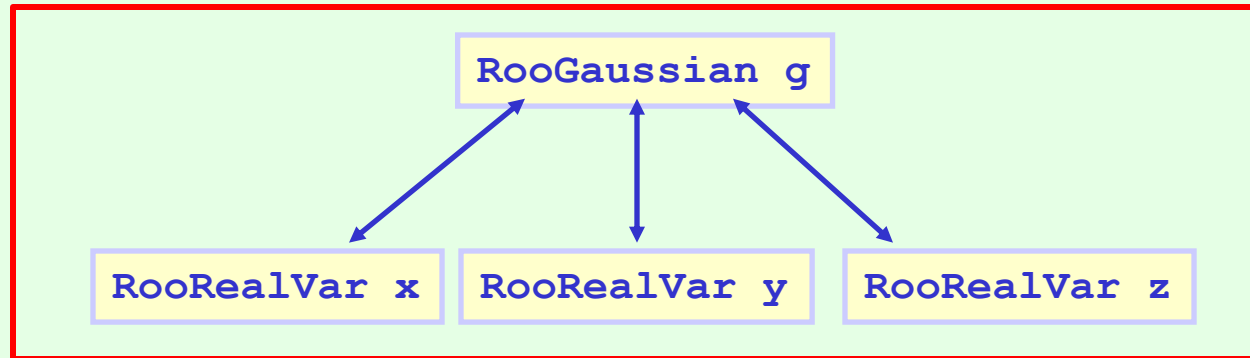
- The workspace serves a container class for all objects created

Math

Gauss( $x, \mu, \sigma$ )

RooWorkspace

RooFit  
diagram



RooFit  
code

```
RooRealVar x("x","x",-10,10) ;  
RooRealVar m("m","y",0,-10,10) ;  
RooRealVar s("s","z",3,0.1,10) ;  
RooGaussian g("g","g",x,m,s) ;  
RooWorkspace w("w") ;  
w.import(g) ;
```

# Populating a workspace the easy way – “the factory”

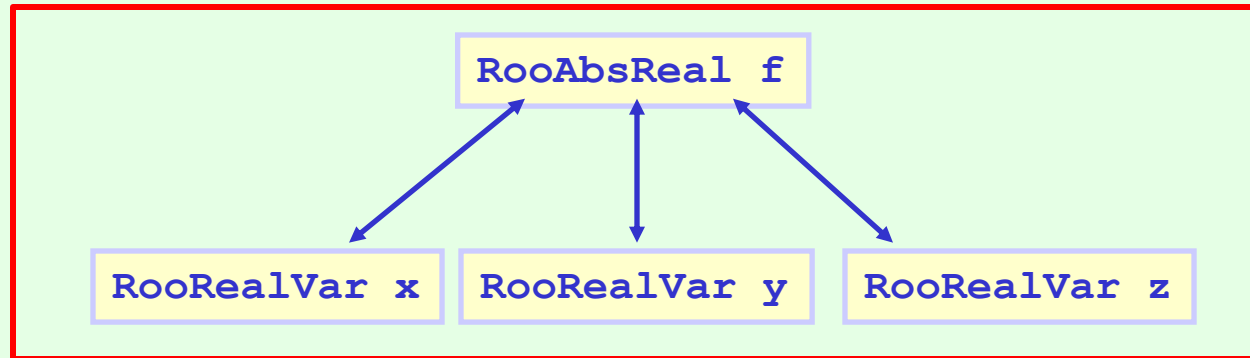
- The **factory** allows to fill a workspace with pdfs and variables using a simplified scripting language

Math

Gauss( $x, \mu, \sigma$ )

RooWorkspace

RooFit  
diagram

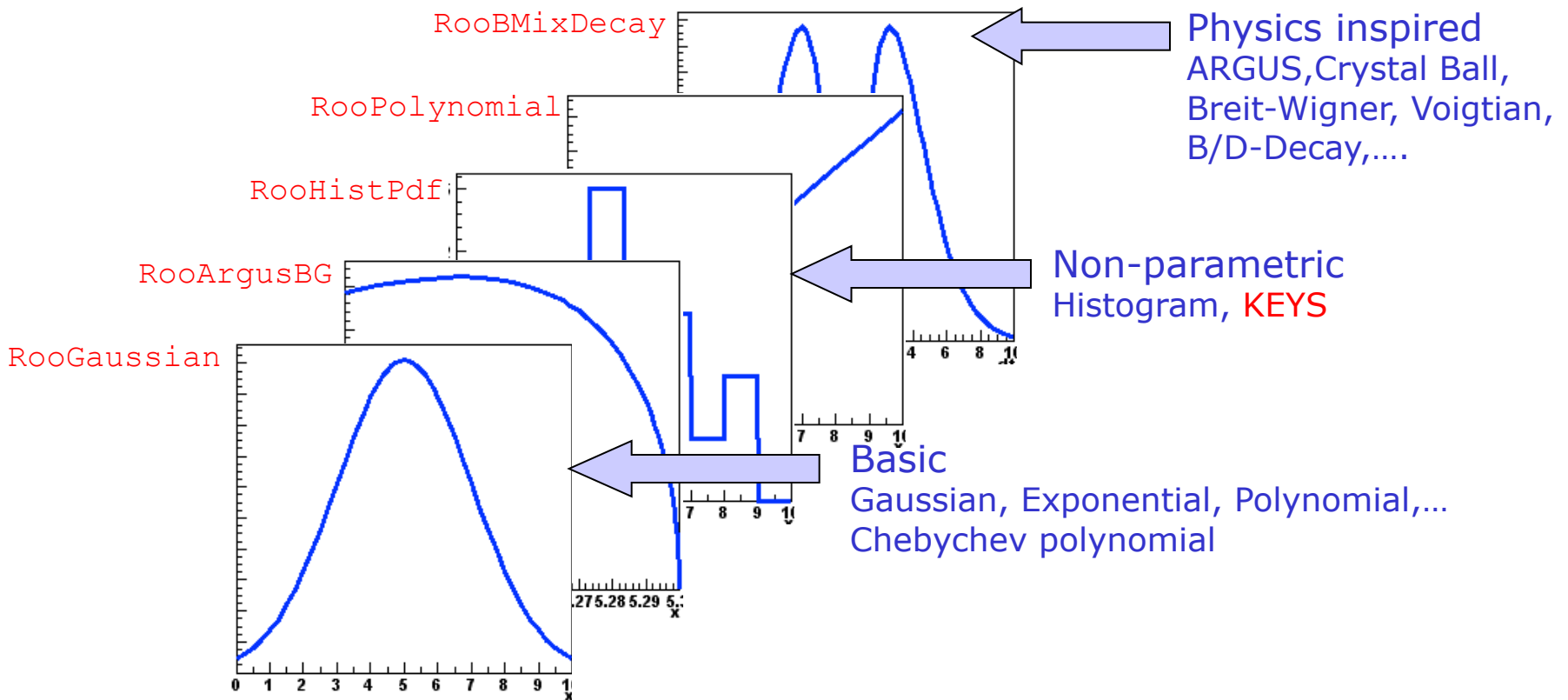


RooFit  
code

```
RooWorkspace w("w") ;  
w.factory("RooGaussian::g(x[-10,10],m[-10,10],z[3,0.1,10])") ;
```

# Model building – (Re)using standard components

- RooFit provides a collection of compiled standard PDF classes



Easy to extend the library: each p.d.f. is a separate C++ class

## Model building – (Re)using standard components

- List of most frequently used pdfs and their factory spec

Gaussian

**Gaussian::g(x, mean, sigma)**

Breit-Wigner

**BreitWigner::bw(x, mean, gamma)**

Landau

**Landau::l(x, mean, sigma)**

Exponential

**Exponential::e(x, alpha)**

Polynomial

**Polynomial::p(x, {a0, a1, a2})**

Chebyshev

**Chebyshev::p(x, {a0, a1, a2})**

Kernel Estimation

**KeysPdf::k(x, dataSet)**

Poisson

**Poisson::p(x, mu)**

Voigtian

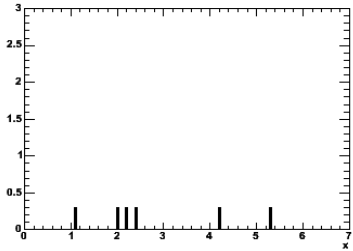
**Voigtian::v(x, mean, gamma, sigma)**

(=BW $\otimes$ G)

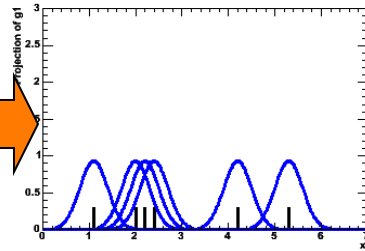
# The power of pdf as building blocks – Advanced algorithms

- Example: a ‘kernel estimation probability model’
  - Construct smooth pdf from unbinned data, using kernel estimation technique

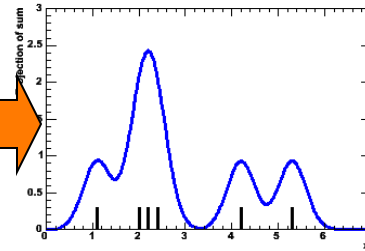
Sample of events



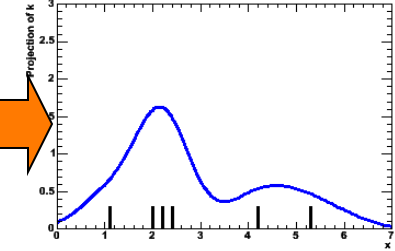
Gaussian pdf for each event



Summed pdf for all events



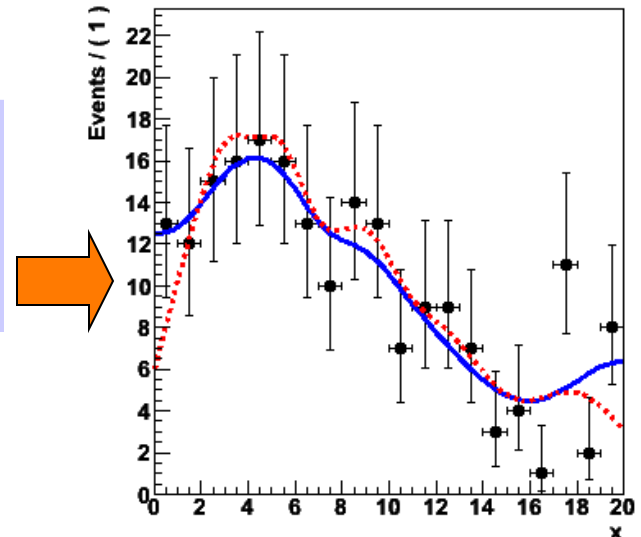
Adaptive Kernel:  
width of Gaussian depends  
on local event density



- Example

```
w. import (myData, Rename ("myData")) ;  
w. factory ("KeysPdf::k(x, myData)") ;
```

- Also available for n-D data



## The power of pdf as building blocks – adaptability

- RooFit pdf classes do not require their parameter arguments to be variables, one can plug in functions as well
- Allows trivial customization, extension of probability models

class RooGaussian

$$Gauss(x | \mu, \sigma)$$

also class RooGaussian!

$$Gauss(x | \underbrace{\mu \cdot (1 + 2\alpha)}_{\text{response function}}, \sigma)$$

Introduce a response function for a systematic uncertainty

```
// Original Gaussian
w.factory("Gaussian::g1(x[80,100],m[91,80,100],s[1])")

// Gaussian with response model in mean
w.factory("expr::m_response(\"m*(1+2alpha)\",m,alpha[-5,5])") ;
w.factory("Gaussian::g1(x,m_response,s[1])")
```

NB: “expr” operates builds an interpreted function expression on the fly

# The power of building blocks – operator expressions

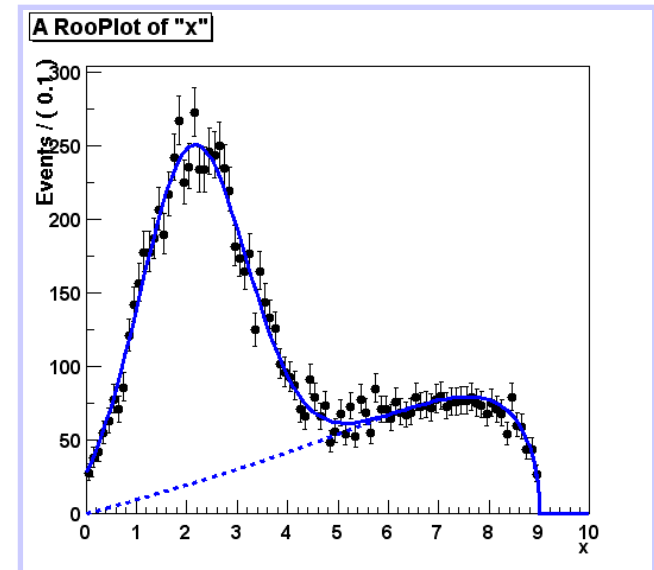
- Create a SUM expression to represent a sum of probability models

```
w.factory("Gaussian::gauss1(x[0,10],mean1[2],sigma[1])" );
w.factory("Gaussian::gauss2(x,mean2[3],sigma)") ;
w.factory("ArgusBG::argus(x,k[-1],9.0)") ;

w.factory("SUM::sum(g1frac[0.5]*gauss1, g2frac[0.1]*gauss2, argus)")
```

- In composite model visualization components can be accessed by name

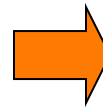
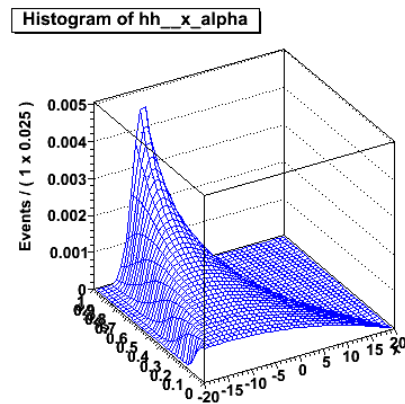
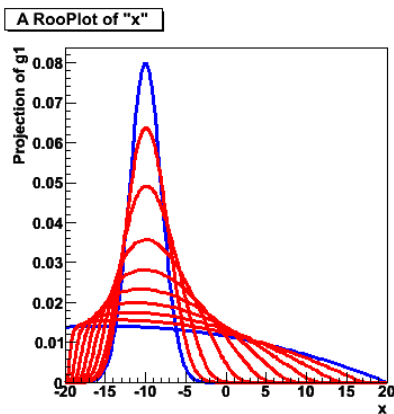
```
// Plot only argus components
w::sum.plotOn(frame,Components("argus"),
              LineStyle(kDashed)) ;
```



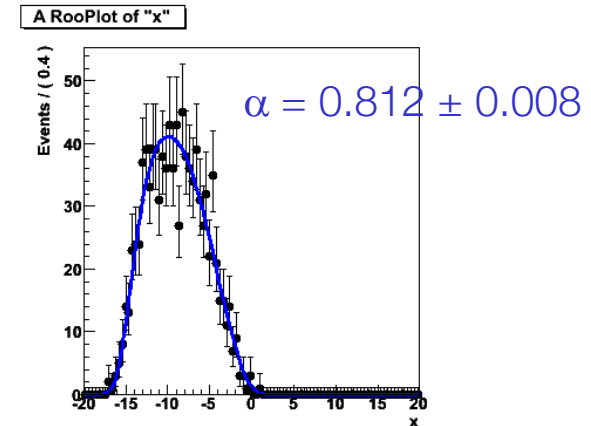
# Powerful operators – Morphing interpolation

- Special operator pdfs can interpolate existing pdf shapes
  - Ex: interpolation between Gaussian and Polynomial

```
w.factory("Gaussian::g(x[-20,20],-10,2)");  
w.factory("Polynomial::p(x,{-0.03,-0.001})");  
w.factory("IntegralMorph::gp(g,p,x,alpha[0,1])");
```



Fit to data



- Three morphing operator classes available
  - `IntegralMorph` (Alex Read).
  - `MomentMorph` (Max Baak).
  - `PiecewiseInterpolation` (via HistFactory)



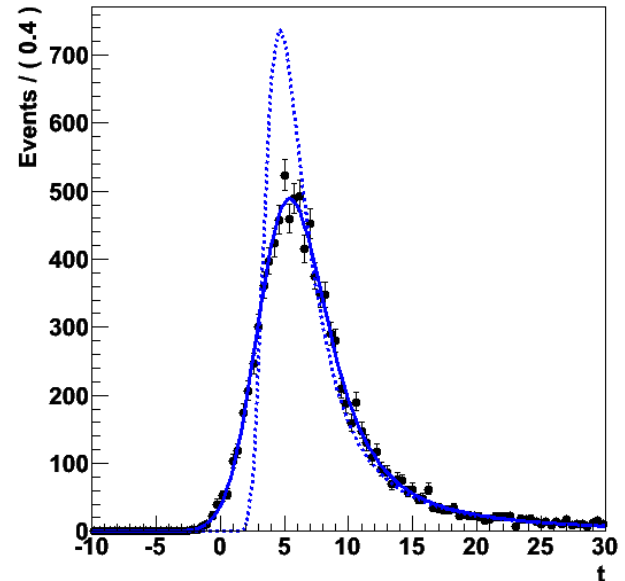
# Powerful operators – Fourier convolution

- Convolve any two arbitrary pdfs with a 1-line expression

```
w.factory("Landau::L(x[-10,30],5,1)") :  
w.factory("Gaussian::G(x,0,2)") ;  
  
w::x.setBins("cache",10000) ; // FFT sampling density  
w.factory("FCONV::LGf(x,L,G)") ; // FFT convolution
```

- Exploits power of FFTW package available via ROOT
  - Hand-tuned assembler code for time-critical parts
  - Amazingly fast: unbinned ML fit to 10.000 events take ~5 seconds!

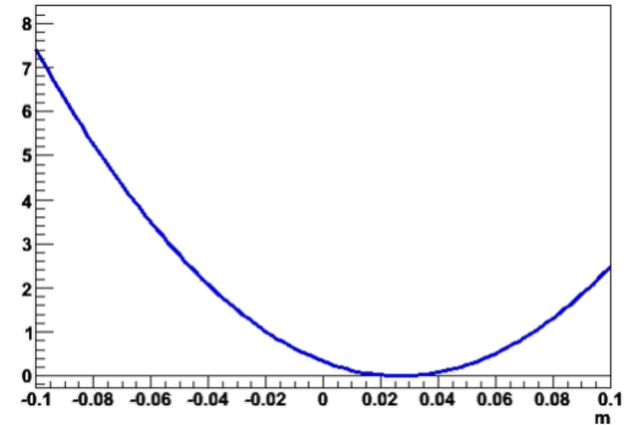
landau (x) gauss convolution



# Working with the likelihood function

- Plot the likelihood function versus a parameter

```
RooAbsReal* nll = w::model.createNLL(data) ;  
  
RooPlot* frame = w::param.frame() ;  
nll->plotOn(frame, ShiftToZero()) ;
```



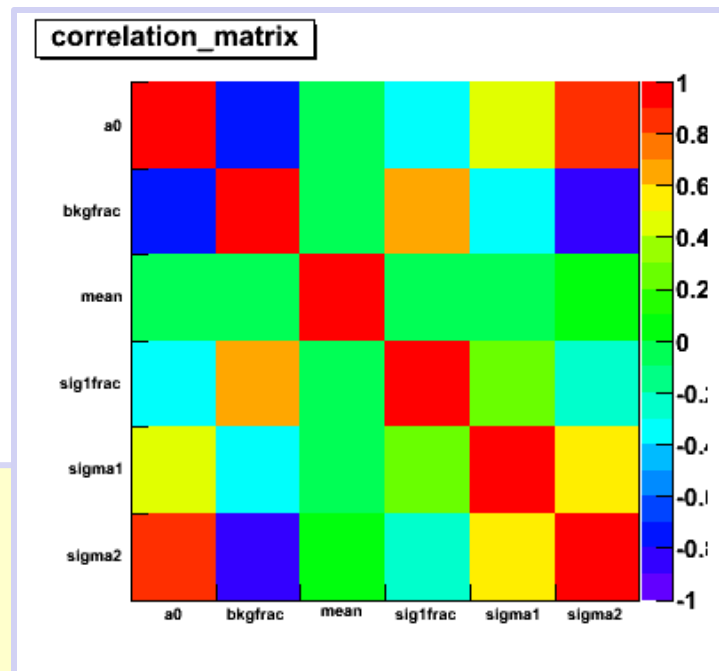
- Maximum Likelihood estimation of parameters and variance

```
RooMinimizer m(*nll) ;  
  
// ML Parameter estimation  
m.minimize("Minuit2","migrad") ;  
  
// Variance estimation  
m.hesse() ;  
  
// Alternatively - all this in one line  
pdf->fitTo(*data) ;
```

# Working with covariance and correlation matrices

- Detailed information on parameter and covariance estimates can be saved for detailed information

```
RoofMinimizer m(*nll) ;  
m.minimize("Minuit2","migrad") ;  
m.hesse() ;  
RooFitResult* r = m.save() ;  
  
// Visualize correlation matrix  
r->correlationHist->Draw("colz") ;  
  
// Extract correlation,covariance matrix  
TMatrixDSym cov = fr->covarianceMatrix() ;  
TMatrixDSym cov = fr->covarianceMatrix(a,b) ;
```



# Use covariance matrices for correlated error propagation

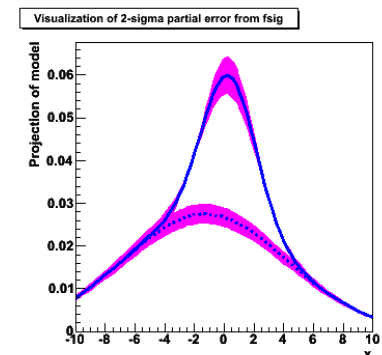
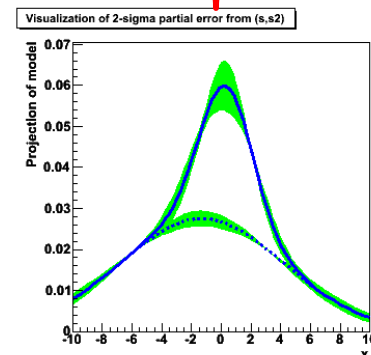
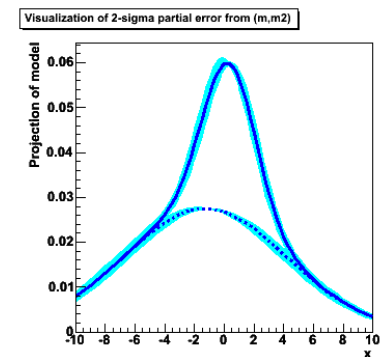
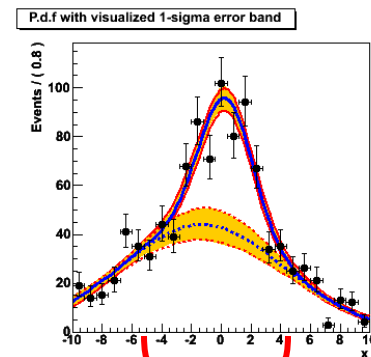
- Can (as visual aid) propagate errors in covariance matrix of a fit result to a pdf projection

```
w::model.plotOn(frame, VisualizeError(*fitresult)) ;  
w::model.plotOn(frame, VisualizeError(*fitresult, fsig)) ;
```

- Linear propagation on pdf projection  $\Delta = \vec{E}V^{-1}\vec{E}$

- Propagated error can be calculated on arbitrary function
  - E.g fraction of events in signal range

```
RooAbsReal* fracSigRange =  
    w::model.createIntegral(x,x,"sig") ;  
  
Double_t err =  
    fracSigRange.getPropagatedError(*fr) ;
```



# Some RooFit practical examples – from start to end

- Signal + Background (analytical)

```
Rooworkspace w("w") ;
```

```
// Construct exponential background model  
w.factory("Exponential::bkg(x[10,100],alpha[-0.04,-0.1,-0])") ;
```

```
// Construct Gaussian signal model  
w.factory("Gaussian::sig(x,mean[40],width[3])") ;
```

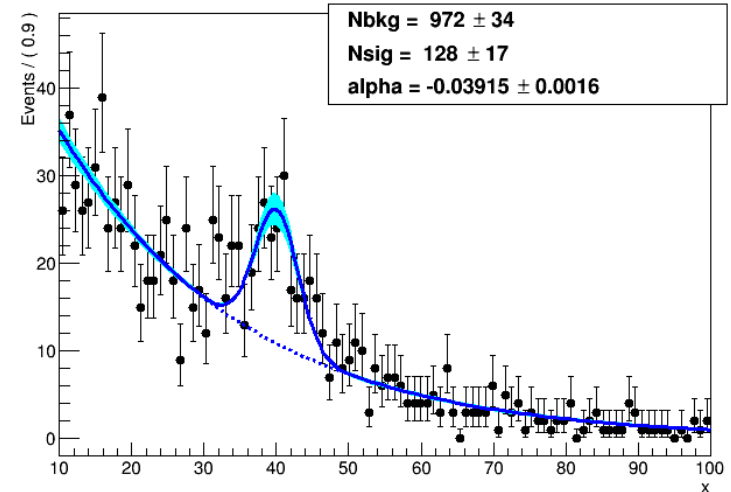
```
// Construct extended ML model of sum of signal and background  
w.factory("SUM::modelsum(Nsig[100,0,200]*sig,Nbkg[1000,0,2000]*bkg)") ;
```

```
// Generate a toy dataset (unbinned) from model, data sample size obtained from expected event count  
RootDataSet* d = w.pdf("modelsum")->generate(*w.var("x")) ;
```

```
// Fit model to toy data  
RootFitResult* r3 = w.pdf("modelsum")->fitTo(*d,Save()) ;
```

```
// Plot data  
RootPlot* frame = w.var("x")->frame() ;  
d->plotOn(frame) ;
```

```
// Plot model (background component separately) and visualization of uncertainties from fit  
w.pdf("modelsum")->plotOn(frame,VisualizeError(*r3)) ;  
w.pdf("modelsum")->plotOn(frame) ;  
w.pdf("modelsum")->plotOn(frame,Components("bkg"),LineStyle(kDashed)) ;  
w.pdf("modelsum")->paramOn(frame) ;  
frame->Draw() ;
```



# Some RooFit practical examples – from start to end

- Two-dimensional signal:  $f(x|y)*g(y)$

```
RoofWorkspace w("w") ;

// Construct  $g(x|y,0.5)$  where the mean of the gaussian
// is a polynomial  $fy=a0+a1*y$ 
w.factory("PolyVar::fy(y[-5,5],{a0[-0.5,-5,5],a1[-0.5,-1,1]})") ;
w.factory("Gaussian::gx(x[-5,5],fy,sigmax[0.5])") ;

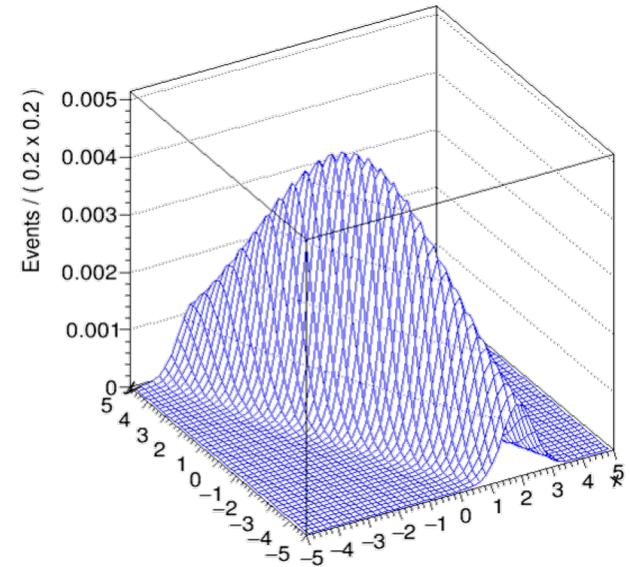
// Construct  $g(y)$ 
w.factory("Gaussian::gy(y,0,3)") ;

// Construct the conditional product  $g(x|y)*g(y)$ 
w.factory("PROD::model(gx|y,gy)") ;

// Generate 1000 events in x and y from model
RootDataSet *data = w.pdf("model")->generate(RooArgSet(*w.var("x"),*w.var("y")),10000) ;

// Plot x distribution of data and projection of model on  $x = \int(dy) \text{model}(x,y)$ 
RootPlot* xframe = w.var("x")->frame() ;
data->plotOn(xframe) ;
w.pdf("model")->plotOn(xframe) ;

// Make two-dimensional plot in x vs y
TH1* hh_model = w.pdf("model")->createHistogram("hh_model",*w.var("x"),Binning(50),
YVar(*w.var("y"),Binning(50))) ;
hh_model->SetLineColor(kBlue) ;
```



# Some RooFit practical examples – from start to end

- **Signal + Background (templates)**

Method 1: Construct unit-normalized pdf from histograms

Model parameters are absolute event counts

```
RoofitWorkspace w("w") ;
```

```
// Import template histograms into workspace
```

```
w.import(*h_bkg,Rename("histo_bkg")) ;
```

```
w.import(*h_sig,Rename("histo_sig")) ;
```

```
// Construct sum of histogram-shaped templates
```

```
w.factory("SUM::modelsum(Nsig[100,0,200]*HistPdf::sig(x[10,100],histo_sig),  
          Nbkg[1000,0,2000]*HistPdf::bkg(x,histo_bkg))") ;
```

```
// Generate a toy dataset (unbinned) from model, data sample size obtained from expected event count
```

```
RootDataSet* d = w.pdf("modelsum")->generate(*w.var("x")) ;
```

```
// Fit model to toy data
```

```
RootFitResult* r3 = w.pdf("modelsum")->fitTo(*d,Save()) ;
```

```
// Plot data
```

```
RootPlot* frame = w.var("x")->frame() ;
```

```
d->plotOn(frame) ;
```

```
// Plot model (background component separately) and visualization of uncertainties from fit
```

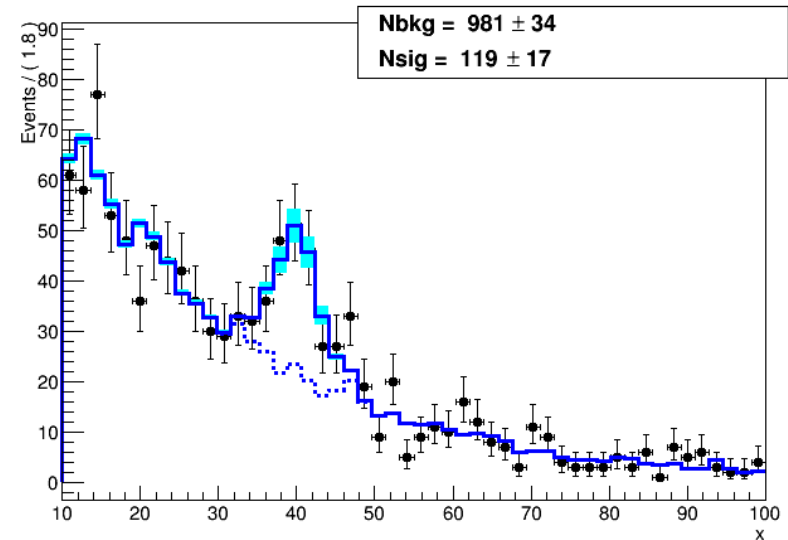
```
w.pdf("modelsum")->plotOn(frame,VisualizeError(*r3)) ;
```

```
w.pdf("modelsum")->plotOn(frame) ;
```

```
w.pdf("modelsum")->plotOn(frame,Components("bkg"),LineStyle(kDashed)) ;
```

```
w.pdf("modelsum")->paramOn(frame) ;
```

```
frame->Draw() ;
```



# Some RooFit practical examples – from start to end

- **Signal + Background (templates)**  
Method 2: Construct event-count scaled pdf from histograms  
Model parameters are scale factors relative histogram counts

```
RoofitWorkspace w("w") ;
```

```
// Import template histograms into workspace
```

```
w.import(*h_bkg,Rename("histo_bkg")) ;
```

```
w.import(*h_sig,Rename("histo_sig")) ;
```

```
// Construct sum of histogram-shaped templates
```

```
w.factory("ASUM::modelsum(kappa_sig[0.01,-0.1,1]*HistFunc::sig(x[10,100],histo_sig),  
kappa_bkg[0.1,-0.1,1]*HistFunc::bkg(x,histo_bkg))") ;
```

```
// Generate a toy dataset (unbinned) from model, data sample size obtained from expected event count
```

```
RootDataSet* d = w.pdf("modelsum")->generate(*w.var("x")) ;
```

```
// Fit model to toy data
```

```
RootFitResult* r3 = w.pdf("modelsum")->fitTo(*d,Save()) ;
```

```
// Plot data
```

```
RootPlot* frame = w.var("x")->frame() ;
```

```
d->plotOn(frame) ;
```

```
// Plot model (background component separately) and visualization of uncertainties from fit
```

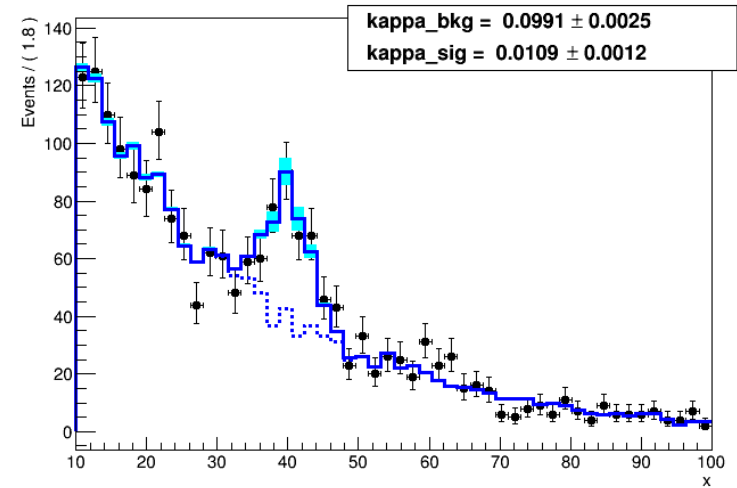
```
w.pdf("modelsum")->plotOn(frame,VisualizeError(*r3)) ;
```

```
w.pdf("modelsum")->plotOn(frame) ;
```

```
w.pdf("modelsum")->plotOn(frame,Components("bkg"),LineStyle(kDashed)) ;
```

```
w.pdf("modelsum")->paramOn(frame) ;
```

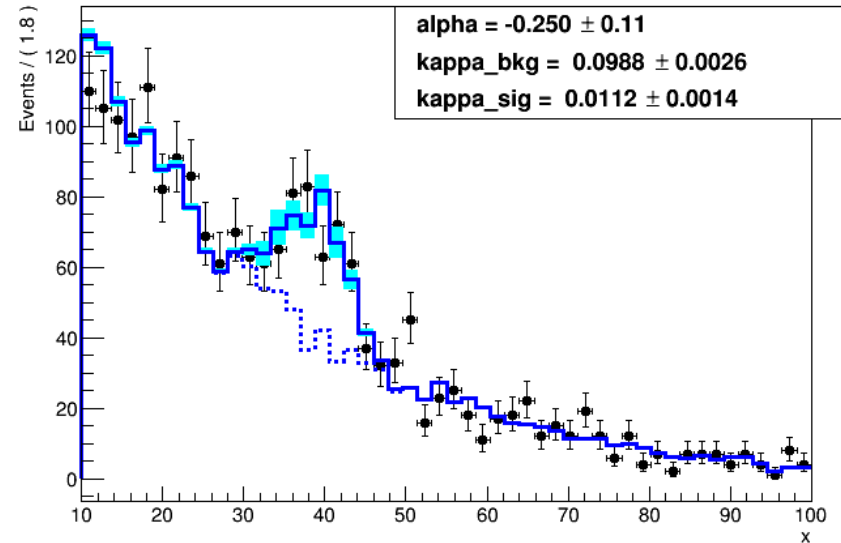
```
frame->Draw() ;
```





# Some RooFit practical examples – from start to end

- Signal + Background (templates)  
With morphing shape parameter



```
RoofitWorkspace w("w") ;
```

```
// Import template histograms into workspace  
w.import(*h_bkg,Rename("histo_bkg")) ;  
w.import(*h_sig_lo,Rename("histo_sig_lo")) ;  
w.import(*h_sig_nom,Rename("histo_sig_nom")) ;  
w.import(*h_sig_hi,Rename("histo_sig_hi")) ;
```

```
w.factory("PiecewiseInterpolation::sig_morph(HistFunc::sig_nom(x,histo_sig_nom),  
HistFunc::sig_lo(x,histo_sig_lo),  
HistFunc::sig_hi(x,histo_sig_hi),alpha[-5,5])") ;
```

```
// Construct sum of histogram-shaped templates  
w.factory("ASUM::modelsum(kappa_sig[0.01,-0.1,1]*sig_morph,  
kappa_bkg[0.1,-0.1,1]*HistFunc::bkg(x,histo_bkg))") ;
```

```
// Generate a toy dataset (unbinned) from model, data sample size obtained from expected event count  
RooDataSet* d = w.pdf("modelsum")->generate(*w.var("x")) ;
```

```
// Fit model to toy data  
RooFitResult* r3 = w.pdf("modelsum")->fitTo(*d,Save()) ;
```

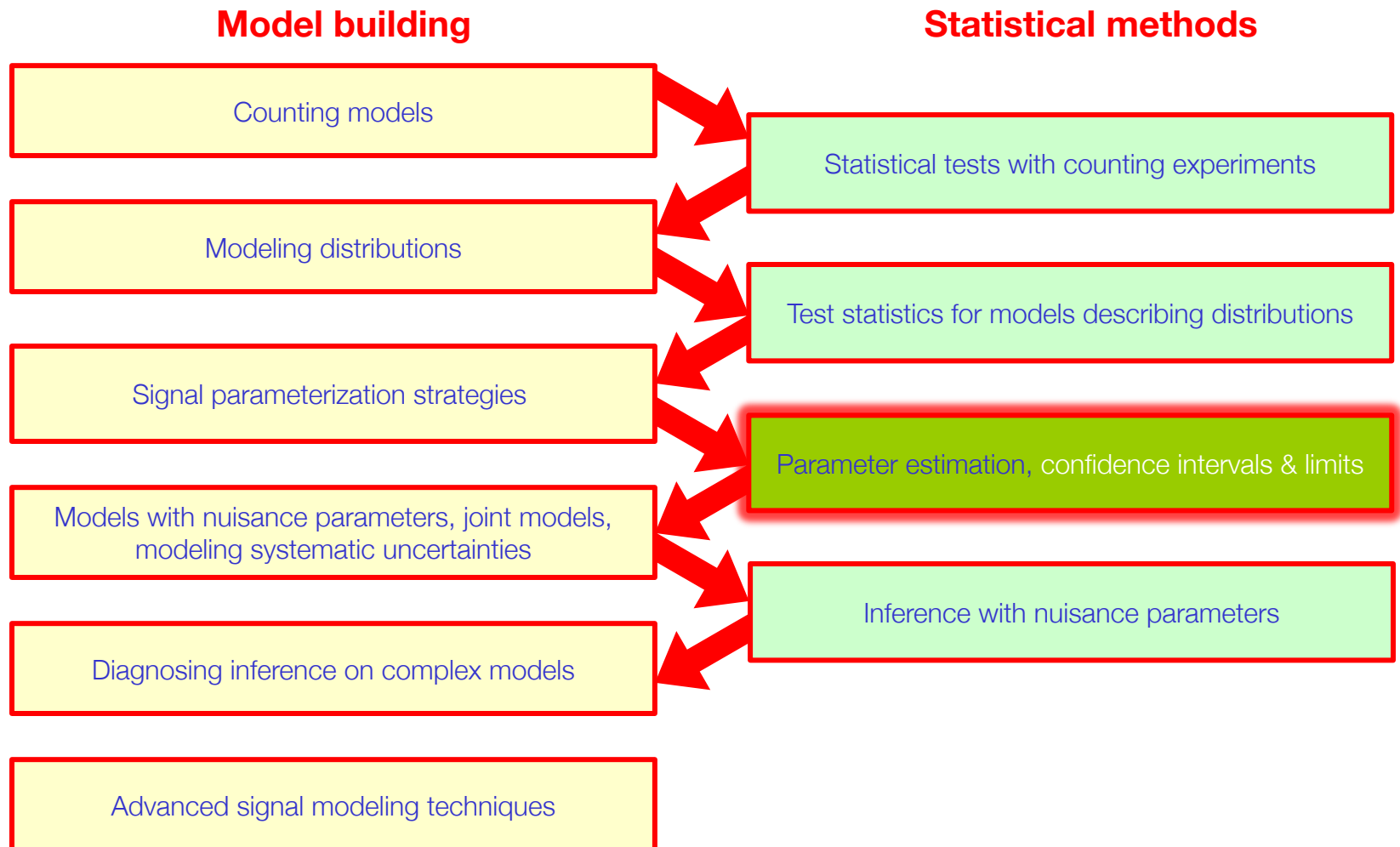
```
// Plot data  
RooPlot* frame = w.var("x")->frame() ;  
d->plotOn(frame) ;
```

# Statistical methods 3

Inference with parameters:  
maximum likelihood, confidence  
intervals, upper limits, likelihood ratio  
and asymptotic formulae

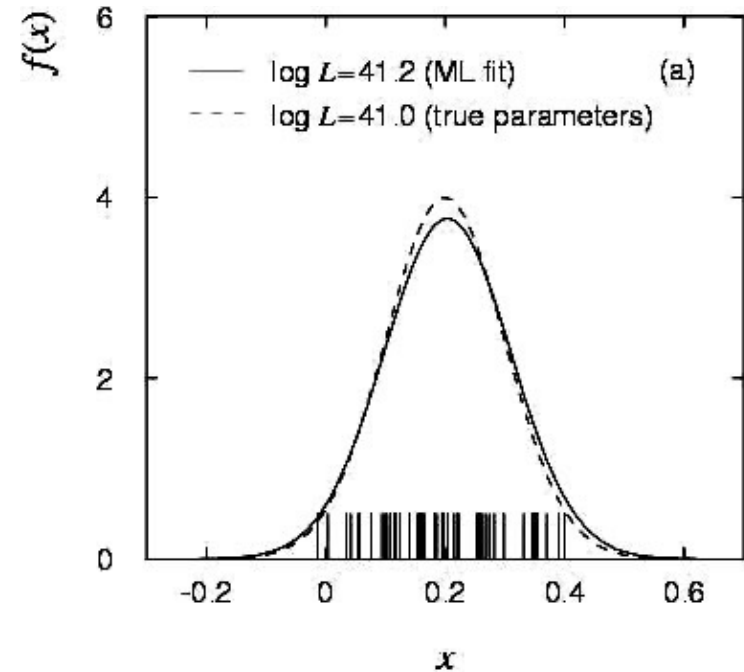
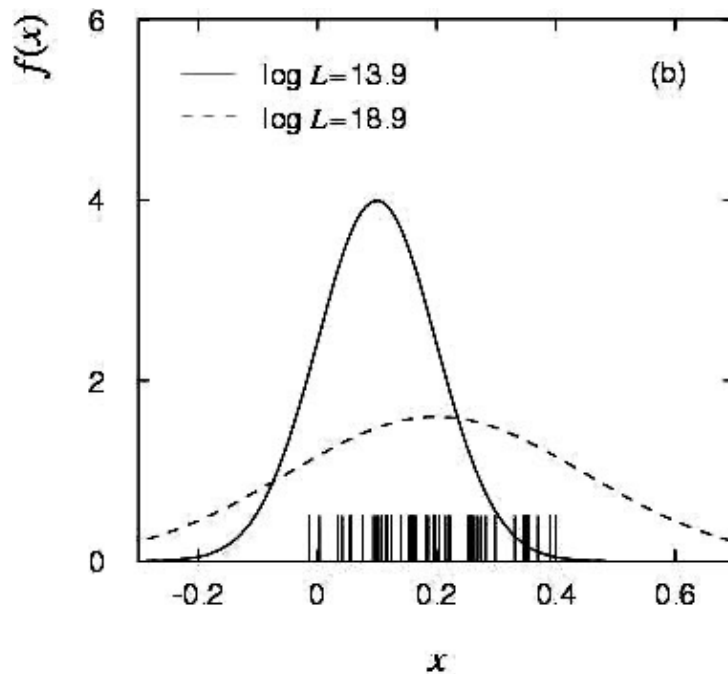
# Roadmap of this course

- Start with basics, gradually build up to complexity



# Parameter estimation using Maximum Likelihood

- Likelihood is high for values of  $p$  that result in distribution similar to data



- Define the **maximum likelihood (ML) estimator** to be the procedure that finds the parameter value for which the likelihood is maximal.

## Parameter estimation – Maximum likelihood

- Practical estimation of maximum likelihood performed by minimizing the negative log-Likelihood

$$L(\vec{p}) = \prod_i f(\vec{x}_i; \vec{p})$$



$$-\ln L(\vec{p}) = -\sum_i \ln F(\vec{x}_i; \vec{p})$$

- Advantage of log-Likelihood is that contributions from events can be summed, rather than multiplied (computationally easier)
- In practice, find point where derivative of  $-\log L$  is zero

$$\left. \frac{d \ln L(\vec{p})}{d \vec{p}} \right|_{p_i = \hat{p}_i} = 0$$

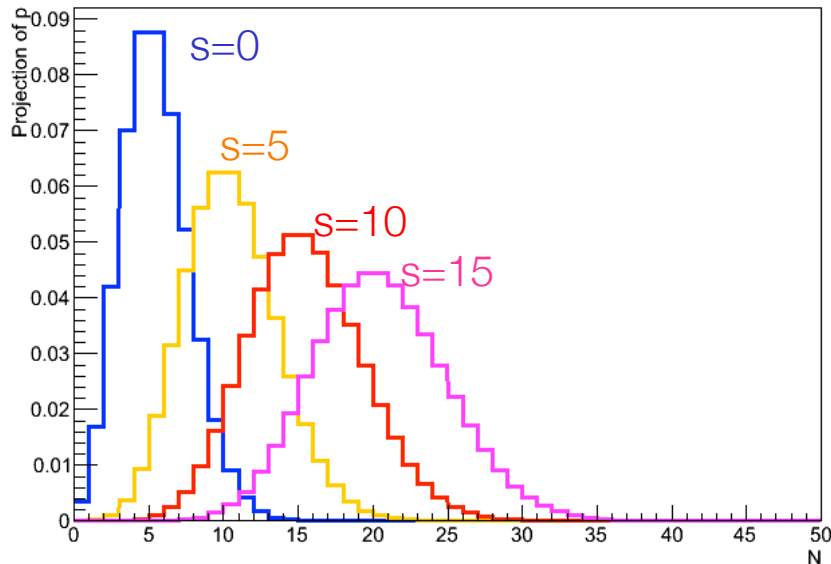
- Standard notation for ML estimation of  $p$  is  $\hat{p}$

# Example of Maximum Likelihood estimation

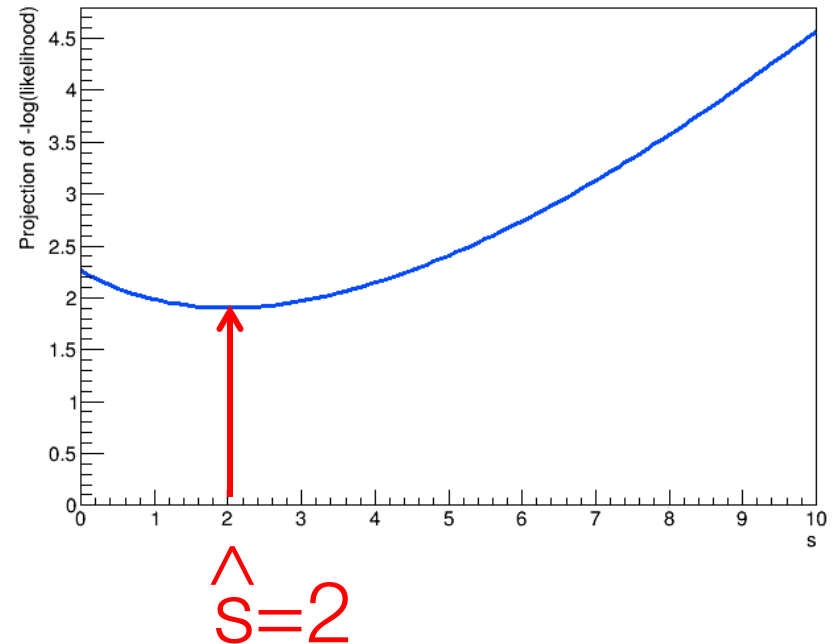
- Illustration of ML estimate on Poisson counting model

$$L(N | s) = \text{Poisson}(N | s + \tilde{b})$$

$-\log L(N|s)$  versus  $N$  [ $s=0,5,10,15$ ]



$-\log L(N|s)$  versus  $s$  [ $N=7$ ]



- Note that Poisson model is discrete in  $N$ , *but continuous in  $s$ !*

# Properties of Maximum Likelihood estimators

- In general, Maximum Likelihood estimators are
  - Consistent (gives right answer for  $N \rightarrow \infty$ )
  - Mostly unbiased (bias  $\propto 1/N$ , may need to worry at small N) `\ex05.C'`
  - Efficient for large N (you get the smallest possible error)
  - Invariant: (a transformation of parameters will Not change your answer, e.g.  $(\hat{p})^2 = \widehat{(p^2)}$ )
- MLE efficiency theorem: *the MLE will be unbiased and efficient if an unbiased efficient estimator exists*
  - Proof not discussed here
  - Of course this does not guarantee that any MLE is unbiased and efficient for any given problem

# Relation between Likelihood and $\chi^2$ estimators

- Properties of  $\chi^2$  estimator follow from properties of ML estimator using *Gaussian probability density functions*

$$F(x_i, y_i, \sigma_i; \vec{p}) = \prod_i \exp \left[ - \left( \frac{y_i - f(x_i; \vec{p})}{\sigma_i} \right)^2 \right]$$

← Gaussian Probability Density Function in  $p$  for single measurement  $y \pm \sigma$  from a predictive function  $f(x|p)$



Take log,  
Sum over all points  $(x_i, y_i, \sigma_i)$

$$-\ln L(\vec{p}) = \frac{1}{2} \sum_i \left( \frac{y_i - f(x_i; \vec{p})}{\sigma_i} \right)^2 = \frac{1}{2} \chi^2$$

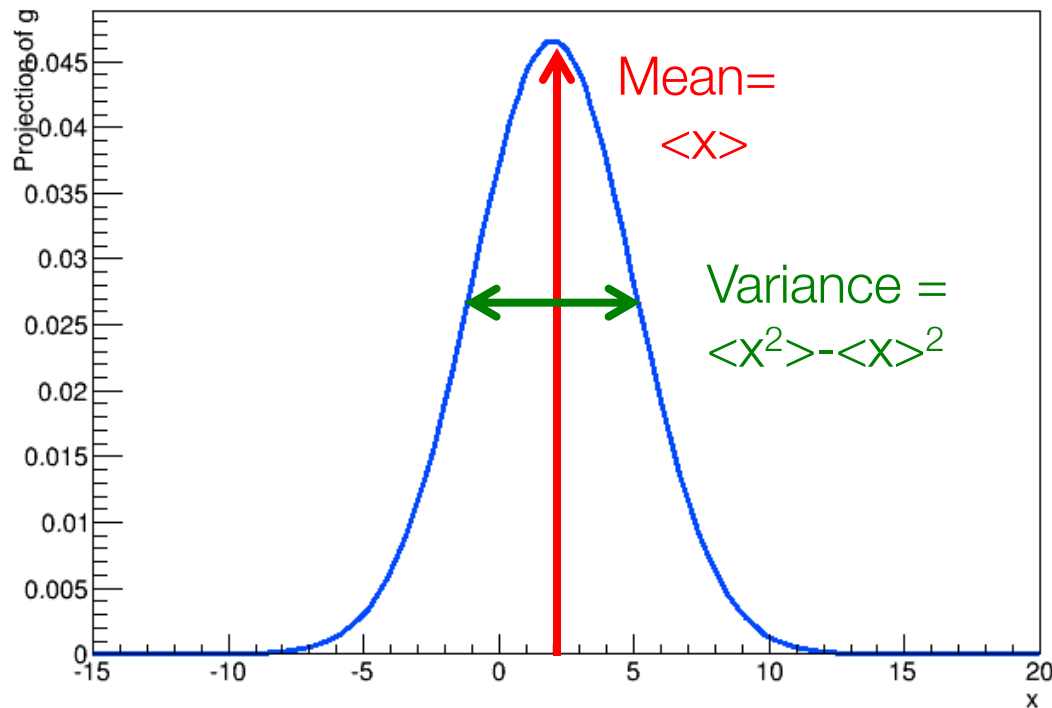
← The Likelihood function in  $p$  for given points  $x_i(s_i)$  and function  $f(x_i; p)$

- The  $\chi^2$  estimator follows from ML estimator, i.e it is
  - Efficient, consistent, bias  $1/N$ , invariant,
  - But only in the limit that the error **on  $x_i$**  is truly Gaussian



# Estimating parameter variance

- Note that ‘uncertainty’ on a parameter estimate is an ambiguous statement
- Can either mean an **interval with a stated confidence or credible, level (e.g. 68%)**, or simply assume it is the **square-root of the variance** of a distribution



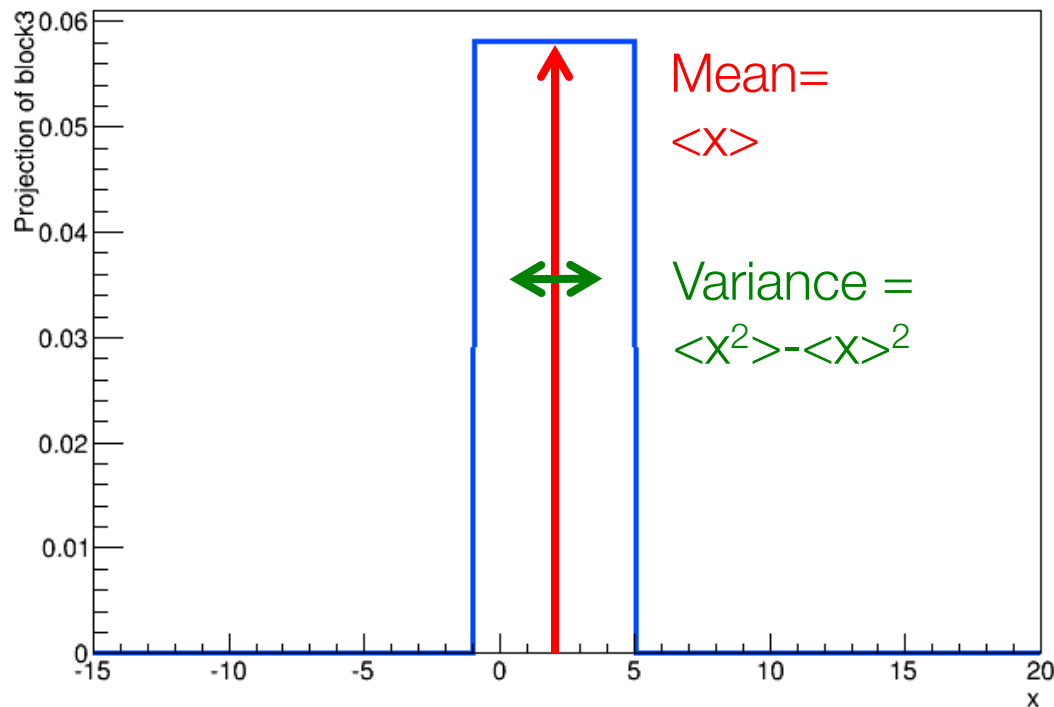
For a Gaussian distribution mean and variance map to parameters for *mean* and *sigma*<sup>2</sup>

and interval defined by  $\sqrt{V}$  contains 68% of the distribution (=‘1 sigma’ by definition)

Thus for Gaussian distributions all common definitions of ‘error’ work out to the same numeric value

# Estimating parameter variance

- Note that 'error' or 'uncertainty' on a parameter estimate is an ambiguous statement
- Can either mean an **interval with a stated confidence or credible, level (e.g. 68%)**, or simply assume it is the **square-root of the variance** of a distribution



For other distributions intervals by  $\sqrt{V}$  do not necessarily contain 68% of the distribution

# Estimating variance on parameters

- Variance on of parameter can also be estimated from Likelihood using the variance estimator

$$\hat{\sigma}(p)^2 = \hat{V}(p) = \left( \frac{d^2 \ln L}{d^2 p} \right)^{-1}$$

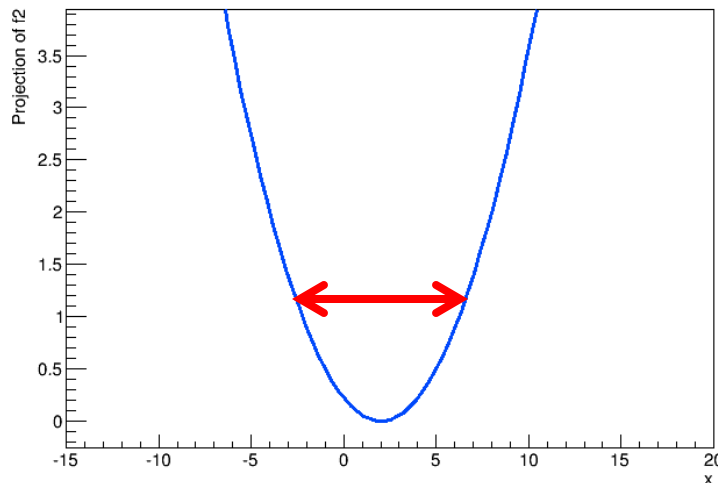
From Rao-Cramer-Frechet inequality

$$V(\hat{p}) \geq 1 + \frac{db}{dp} \bigg/ \left( \frac{d^2 \ln L}{d^2 p} \right)$$

b = bias as function of p,  
inequality becomes equality  
in limit of efficient estimator

- Valid if estimator is **efficient** and **unbiased!**

- Illustration of Likelihood Variance estimate on a Gaussian distribution



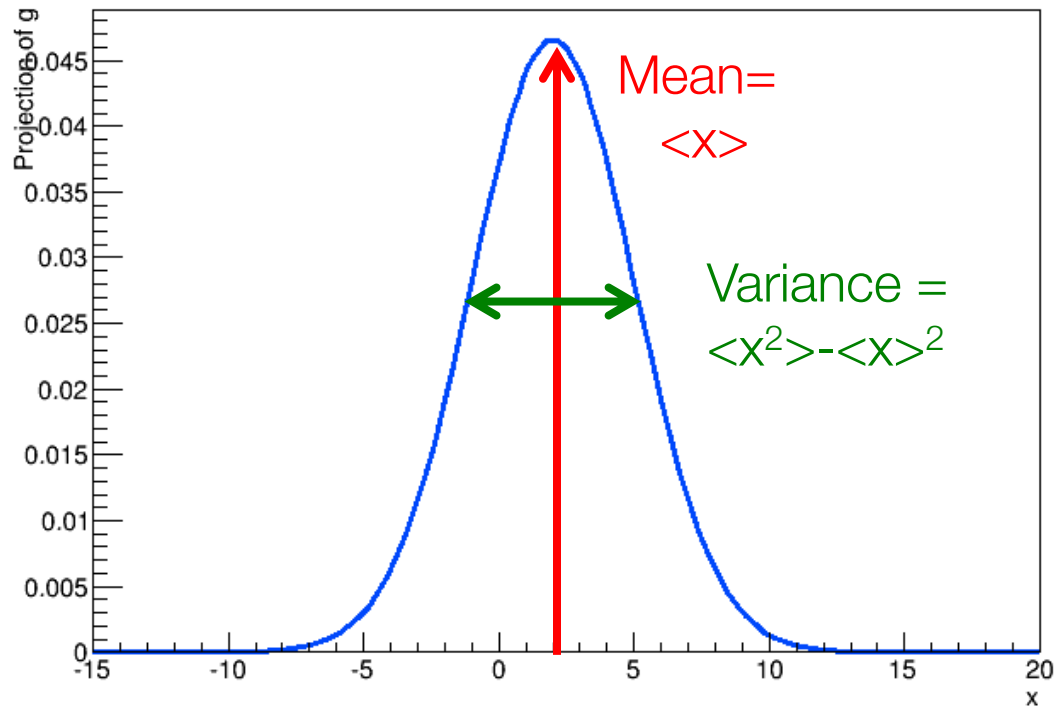
$$f(x | \mu, \sigma) = \frac{1}{\sigma \sqrt{2\pi}} \exp \left[ -\frac{1}{2} \left( \frac{x - \mu}{\sigma} \right)^2 \right]$$

$$\ln f(x | \mu, \sigma) = -\ln \sigma - \ln \sqrt{2\pi} + \frac{1}{2} \left( \frac{x - \mu}{\sigma} \right)^2$$

$$\left. \frac{d \ln f}{d \sigma} \right|_{x=\mu} = \frac{-1}{\sigma} \Rightarrow \left. \frac{d^2 \ln f}{d^2 \sigma} \right|_{x=\mu} = \frac{1}{\sigma^2}$$

# Bayesian parameter estimation

- Bayesian parameter estimate is the posterior mean
- Bayesian variance is the posterior variance

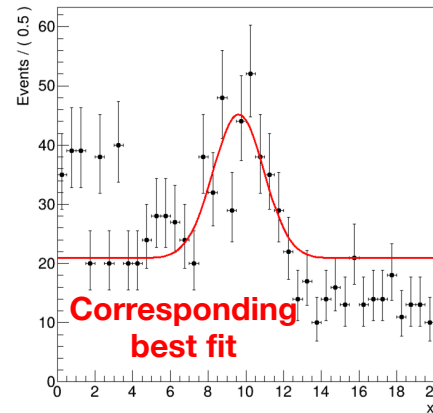
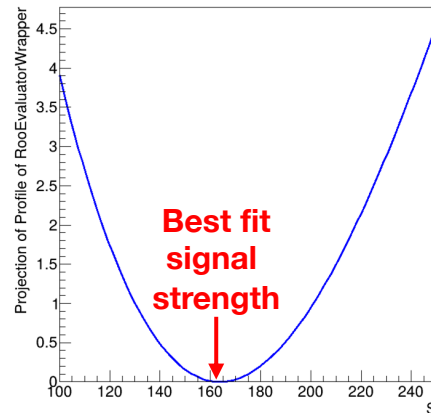


$$\hat{\mu} = \int \mu P(\mu | N) d\mu$$

$$\hat{V} = \int (\hat{\mu} - \mu)^2 P(\mu | N) d\mu$$

# Goodness-of-fit

- An important question that arises in all statistical modeling:  
**is “is the best-fit” actually a “good fit”?**



- Fit only considers degrees of freedom expressed in the likelihood → might not capture observed data
- In other words, the ‘alternative hypotheses’ considered is the ensemble of parameter values of the model
- Can you quantify goodness-of-fit ‘abstractly’, without an explicit alternative hypothesis defining the whole space of possibilities?
  - Generally a hard problem, *no approach is assumption-free*
- Commonly used: reduced  $\chi^2$ .
  - Effectively it is p-value to obtain observed data under hypothesis of best-fit.
  - Implicit alternative hypothesis: independent Gaussian fluctuations in each bin
  - Not always a realistic assumption for deviations (ignores systematic effects)
- Much more on goodness-of-fit tomorrow → Lecture by Lydia

# What can we do with composite hypothesis

- With simple hypotheses – inference is restricted to making statements about  $P(D|\text{hypo})$  or  $P(\text{hypo}|D)$
- With composite hypotheses – many more options

- **1 Parameter estimation and variance estimation**

- What is value of  $s$  for which the observed data is most probable?
  - What is the variance (std deviation squared) in the estimate of  $s$ ?
- }  $s=5.5 \pm 1.3$

- **2 Confidence intervals**

- Statements about model parameters using frequentist concept of probability
- $s < 12.7$  at 95% confidence level
- $4.5 < s < 6.8$  at 68% confidence level

- **3 Bayesian credible intervals**

- Bayesian statements about model parameters
- $s < 12.7$  at 95% credibility